Large population limit for a wide class of individual-based epidemic models

Amaury Lambert & SMILE group











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SMILE : An interdisciplinary group in Paris















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SMILE = Stochastic Models for the Inference of Life Evolution

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Kermack-McKendrick 1927

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A more complex example :



 ...Classes/compartments are structuring variables which are observable or play a role in transmission (health condition, employment category, age...)

Under the assumption of large populations and...

- Lack of memory property of
 - the time spent in each compartment
 - the time between consecutive infections
- …Compartmental models typically lead to systems of ODEs of the type

$$\begin{cases} \dot{E} &= \beta SI - \eta E \\ \dot{I} &= \eta E - \gamma I \\ \dot{S} &= -\beta SI \end{cases}$$

Life cycle of SARS-Cov2...



Ferretti et al Science 2020

Requires a large number of compartments

- Non-exponential durations spent in each compartment and between infections
- Q1. How can we avoid using high-dimensional systems of ODEs (large pop limit)?
- Q2. How can we take advantage of the spontaneous growth to large pop?

A general branching process model

- Susceptibles in excess (branching process assumption)
- Each individual is identified by her age a = time elapsed since her infection
- Each infected individual is characterized by
 - (i) A stochastic life-history process $(X(a); a \ge 0)$ where

X(a) =state of ind at age a

- (ii) A stochastic infection point process ${\cal P}$ encoding the times of transmission of the disease to susceptibles
- Example : the SEIR model is recovered by setting

$$X(a) = \begin{cases} E & \text{if } a \in [0, T_l) \\ l & \text{if } a \in [T_l, T_R) \\ R & \text{if } a \in [T_R, \infty) \end{cases} \text{ and } \mathcal{P}_k = \mathbb{1}_{[T_l, T_R]} \mathcal{P}$$

where *P* is a rate β Poisson point process.

Each newly infected individual draws a pair (X, P) in the same distribution, independently of donor state and age

A word on assumptions

Susceptibles in excess (branching process model)

- No re-infection
- No spatial/social structure (mean-field model)
- Independence of receiver and donor (contact matrix with identical rows)
- Allows for any type of correlation and temporal auto-correlation of (X, \mathcal{P})
- Additional feature : each new infection occurring at time t is successful with probability c(t) (lockdown, masks, social distancing...)

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 $\tau(a) da := \mathbb{E}[\mathcal{P}(da)]$

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Define the empirical measure of the population as

$$\mu_t(da \times \{i\}) := \#\{ \text{ infected ind's at time } t, \text{ in state } i \text{ and of } age \in da \}$$
$$= \sum_k \delta_{t-\sigma_k}(da) \mathbb{1}_{X_k(t-\sigma_k)=i}$$

where σ_k is the birth time of individual k.

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where σ_k is the birth time of individual k.

Also define

$$Y_t(i) := \#\{ \text{ infected ind's at time } t, \text{ in state } i \} = \int_{a=0}^{\infty} \mu_t(da \times \{i\})$$
$$Z_t := \#\{ \text{ infected ind's at time } t \} = \sum_i Y_t(i)$$

A law of large numbers

Start with *N* infected individuals at time 0.

Assume that the ages of the initial individuals $(A_k)_k$ are iid with density g.

▶ The initial life-cycle and infection processes are iid $(\tilde{X}_k, \tilde{\mathcal{P}}_k)_k$, where

$$\widetilde{X}_k(a) = X_k(A_k + a), \quad \widetilde{\mathcal{P}}_k(da) = \mathcal{P}_k(A_k + da).$$

Theorem (Foutel–Rodier et al, 2020)

As $N \to \infty$, for any t, $\frac{1}{N}\mu_t^N(da \times \{i\}) \longrightarrow n(t,a) p(a,i) da$

where $p(a, i) = \mathbb{P}(X(a) = i)$ and n is the solution to the McKendrick-von Foerster PDE

$$\frac{\partial n}{\partial a} + \frac{\partial n}{\partial t} = 0$$

with boundary condition

$$n(t,0) = c(t) \int_0^\infty n(t,a)\tau(a) \, da,$$

and initial condition

n(0,a)=g(a).

A law of large numbers : consequences

- The age structure converges to the solution to a PDE that only depends on \(\tau\)
- The class structure $Y_t^N(i)$ converges to

$$\int_0^\infty n(t,a)\,p(a,i)\,da$$

which only depends on $p(a, i) = \mathbb{P}(X(a) = i)$.

- Structuring by age...
 - decouples transmission and life history
 - decouples time and class structure.
- Drawbacks:
 - requires assuming large initial population size
 - requires specifying age profile g, which is not observable.

Branching process counted with its characteristic

Branching process counted with the random characteristic χ

$$Z_t^{\chi} := \sum_k \chi_k(t - \sigma_k),$$

where $\sigma_k =$ infection time of ind k and sum is taken over all ind's in the course of the epidemic.

Examples

$$\chi_k(t) = \begin{cases} 1 & \text{if } k \text{ is in state } i \text{ at time } t \\ 0 & \text{otherwise.} \end{cases} \implies Z_t^{\chi} = Y_t(i)$$
$$\chi_k(t) = \begin{cases} 1 & \text{if } t \ge a \\ 0 & \text{otherwise.} \end{cases} \implies Z_t^{\chi} = \#\{ \text{ ind's of age} \ge a \text{ at time } t \}.$$

Theorem (Jagers 74, Nerman 81, Jagers & Nerman 84)

Assume $\alpha > 0$. Then there is a single random variable W_{∞} such that for all random characteristics χ simultaneously, with probability 1,

$$\lim_{t\to\infty} e^{-\alpha t} Z_t^{\chi} = \frac{\mathbb{E}\left(\widehat{\chi}(\alpha)\right)}{\alpha\beta} W_{\infty},$$

where $\beta := R_0 \int_0^\infty t \, \tau(t) \, e^{-\alpha t} dt$ and

$$\widehat{\chi}(\alpha) := \alpha \int_0^\infty \chi(t) \, e^{-\alpha t} dt,$$

where it is implicit that $\chi(t)$ is the characteristic of the ancestor, 'born' at time 0. In particular, if we take $\chi(t) = \mathbb{1}_{t>a}$, we get

$$\lim_{t \to \infty} \frac{\#\{ \text{ ind's of } age \ge a \text{ at time } t \}}{Z_t} = e^{-\alpha a}$$

A second law of large numbers

- Start with 1 infected individual at time 0.
- Assume $\alpha > 0$.
- ▶ Let T_K be a random time such that $\lim_{K\to\infty} T_K = \infty$, e.g., time when > K infected.
- ► $c(t) \leftarrow c(T_K + t)$.

Theorem (Foutel-Rodier et al, 2020)

Recall Z_t denote the total number of infected by time t. As $K \to \infty$, for any t,

$$\frac{1}{Z_{T_{K}}}\mu_{T_{K}+t}^{K}(da \times \{i\}) \longrightarrow n(t,a)\mathbb{P}(X(a) = i)da \quad a.s.,$$

where n is the solution to the McKendrick-von Foerster PDE

$$\frac{\partial n}{\partial a} + \frac{\partial n}{\partial t} = 0$$

with boundary condition

$$n(t,0) = c(t) \int_0^\infty n(t,a)\tau(a) \, da,$$

and initial condition

$$n(0,a) = \alpha e^{-\alpha a}$$

Application to the COVID-19 epidemic in France

- Given the profile of τ , a compartmental model and a parameter set :
 - α is estimated as the initial growth rate
 - the PDE is solved numerically
 - the one-dimensional marginals $\mathbb{P}(X(a) = i)$ can be computed assuming Gamma distributed sojourn times
 - the observations are assumed to follow a Poisson distribution centered on the predicted value
- The best parameter set is obtained by maximum likelihood



Fitting the incidence curves



Fitting the incidence curves



Discussion

- A general, stochastic, individual-based model for the spread of complex diseases
- Law of large numbers as a result of spontaneous growth, with emerging age structure
- Limit in the form of an age-structured McKendrick-von Foerster PDE decoupling time/class structure and transmission/life history.
- Model is tractable enough to be used for inference.
- Preprint : "From individual-based epidemic models to McKendrick-von Foerster PDEs : A guide to modeling and inferring COVID-19 dynamics", by

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Foutel-Rodier* F, Blanquart F, Courau P, Czuppon P, Duchamps JJ, Gamblin J, Kerdoncuff E, Kulathinal R, Régnier L, Vuduc L, Lambert** A & Schertzer** E (2020).
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* first author, ** co-last authors
    https://arxiv.org/abs/2007.09622
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See also :

JY Fan, K Hamza, P Jagers, F Klebaner. Convergence of the age structure of general schemes of population processes. *Bernoulli* 2020.

K Hamza, P Jagers, F Klebaner. The age structure of population- dependent general branching processes in environments with a high carrying capacity. *Proceedings of the Steklov Institute of Mathematics*, 2013.

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Acknowledgements





...Et merci pour votre attention!

Notation	Description	Value	Source
R ₀	Basic reproduction number during lockdown	0.734	E
W	Total number of infections before March 17 2020	9.52×10^{5}	Е
$p_{ m crit} + p_{ m sev}$	Probability of being hospitalized	0.036	S
$\frac{\rho_{\rm crit}(1-d_{\rm hosp})}{\rho_{\rm crit}+\rho_{\rm sev}}$	Probability of entering ICU conditional on being at the hospital	0.19	S
$\frac{d_{\rm hosp} + (1 - d_{\rm hosp})d_{\rm ICU}}{1 + \rho_{\rm sev}/\rho_{\rm crit}}$	Death probability conditional on being hospitalized	0.181	S
d _{ICU}	Probability of death conditional on being in ICU	0.709	Е
$p_{ m short}$	Probability of a short stay at hospital	0.701	Е
D _{ch}	Delay between severe infection and hospital admission	14.5 days	Е
D_{short}	Delay between hospital admission and quick discharge	7.36 days	E
D_{long}	Delay between hospital admission and slow discharge	47.5 days	Е
Dcu	Delay between critical infection and hospital admission	11.0 days	Е
D _H	Delay between hospital admission and ICU admission	1.5 days	S
D_U	Delay between ICU admission and discharge	28.2 days	Е
D _D	Delay between ICU admission and death	9.90 days	Е
γ	Scale parameter common to all Gamma distributions	0.316	Е

TABLE – Inferred parameter set for the occupancy model. The values indicated for the durations correspond to the means of the Gamma distributions. In the "Source" column, "E" = estimated, "F" = taken from Salje et al (2020).