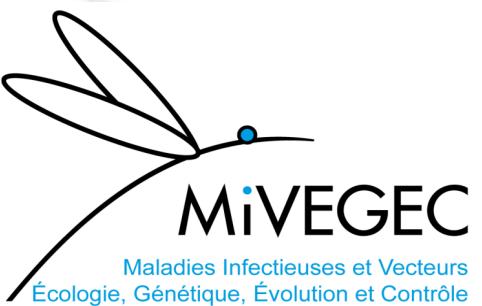


Suivi épidémiologique et contrôle : applications d'une modélisation en temps discret de la dynamique hospitalière de COVID-19 en France

Mircea T. Sofonea, Corentin Boennec, Bastien Reyné, Baptiste Elie, Ramsès Djidjou-Demasse, Christian Selinger, Yannis Michalakis, Samuel Alizon



ETE modelling team,
MIVEGEC (Univ. Montpellier, CNRS, IRD)

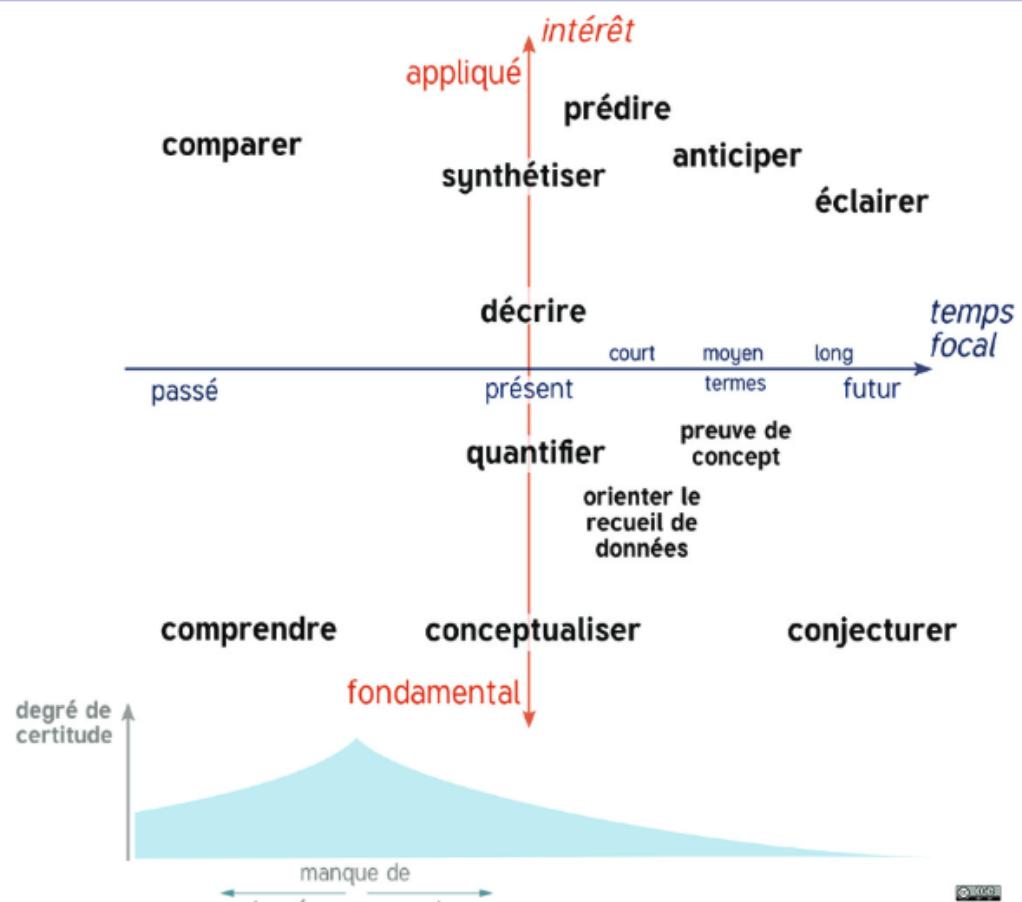


Congrès SMAI 2021,
La Grande Motte 22.VI.2021



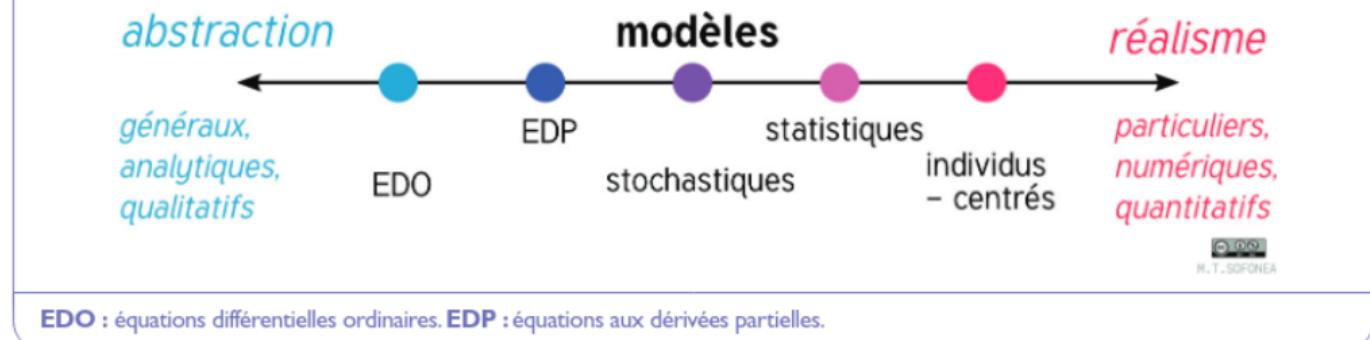
modelling: means and ends

Figure 1. Typologie des questions abordées par les modèles.

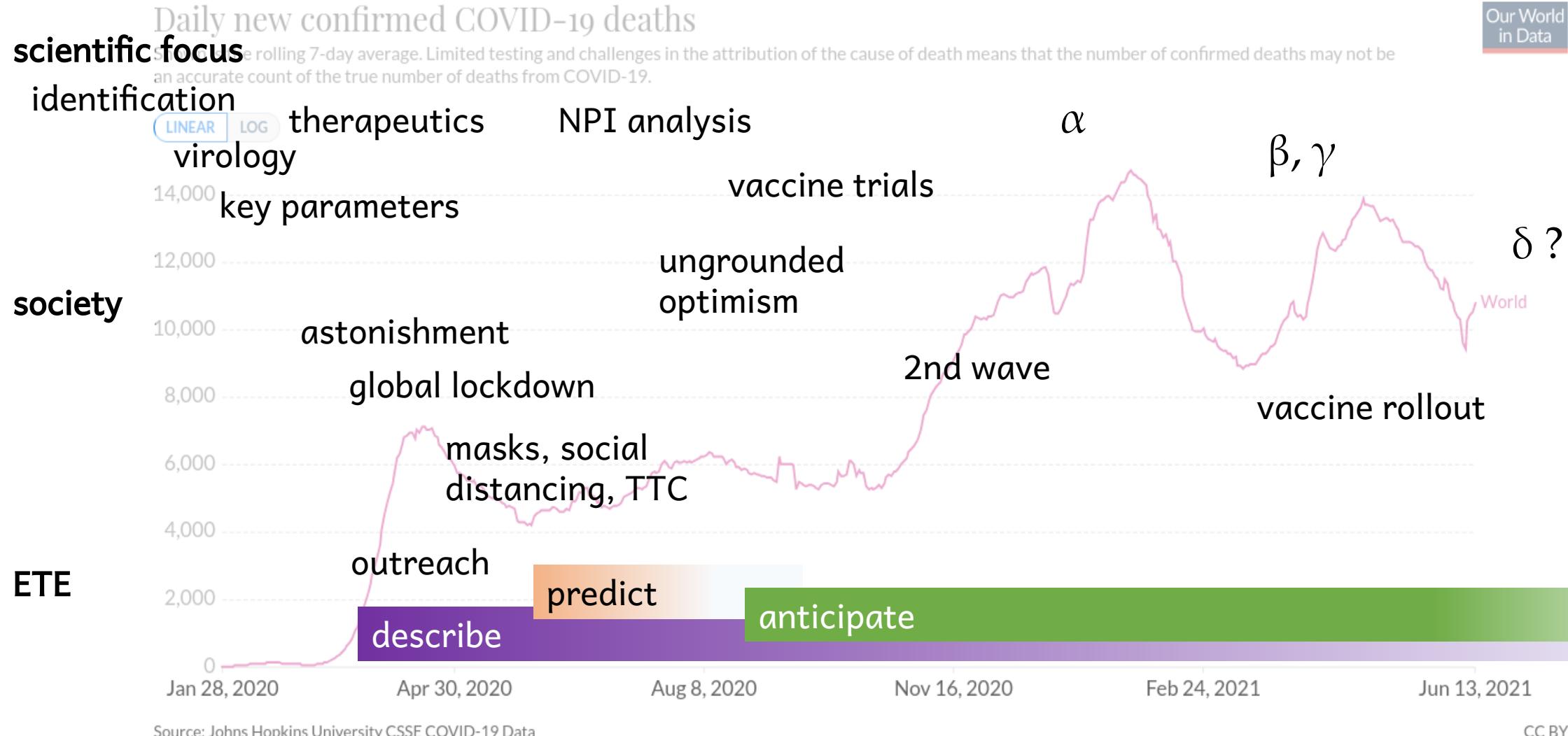


La disposition absolue des objectifs dans le repère ainsi que la distribution du degré de certitude est arbitraire et n'a qu'une valeur indicative.

Figure 2. Principales catégories méthodologiques de modèles rencontrés en épidémiologie mathématiques.



ongoing research on an ongoing pandemic



outline

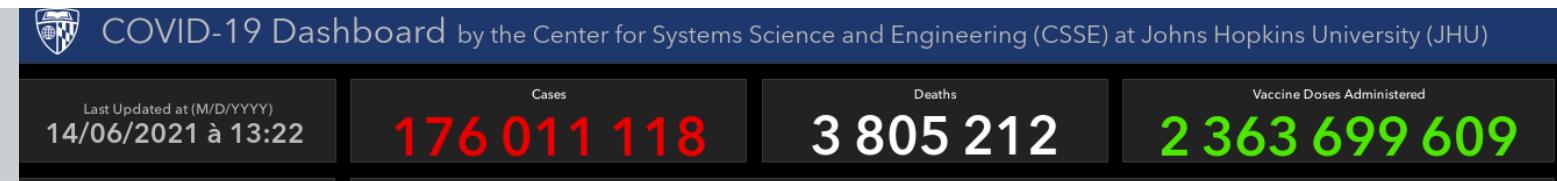
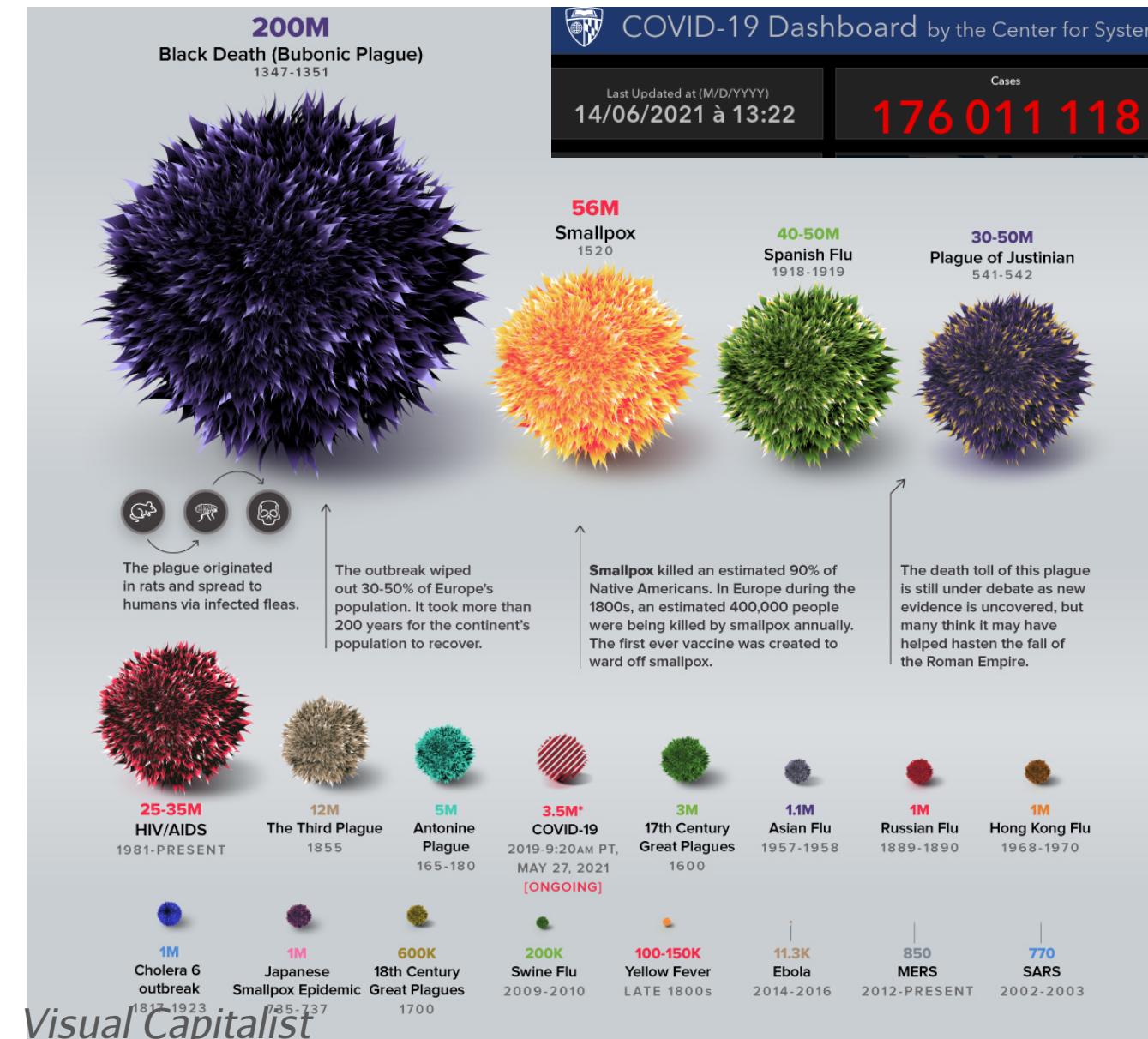
introductory elements

I. Structure

II. Counterfactuals

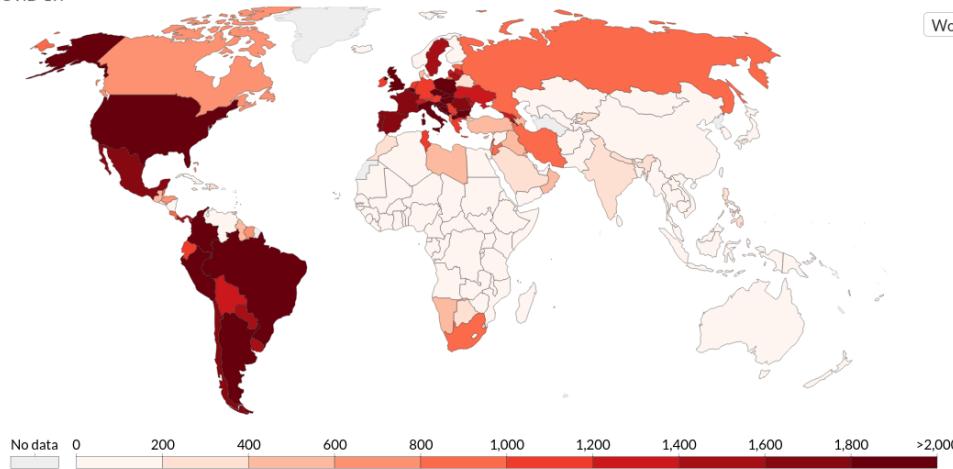
III. Projections

provisional death toll



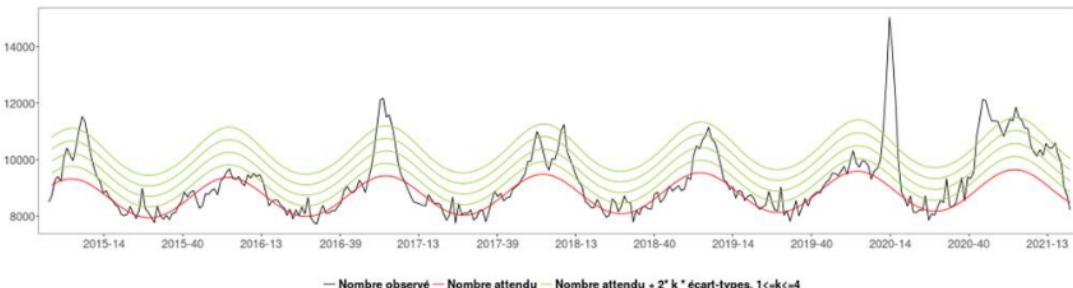
Cumulative confirmed COVID-19 deaths per million people, Jun 15, 2021

Limited testing and challenges in the attribution of the cause of death means that the number of confirmed deaths may not be an accurate count of the true number of deaths from COVID-19.



Our World in Data

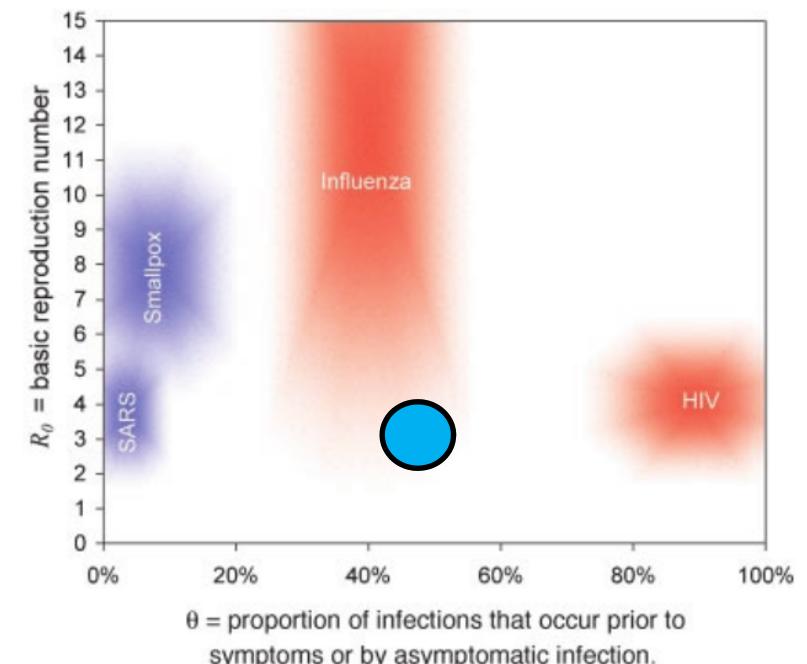
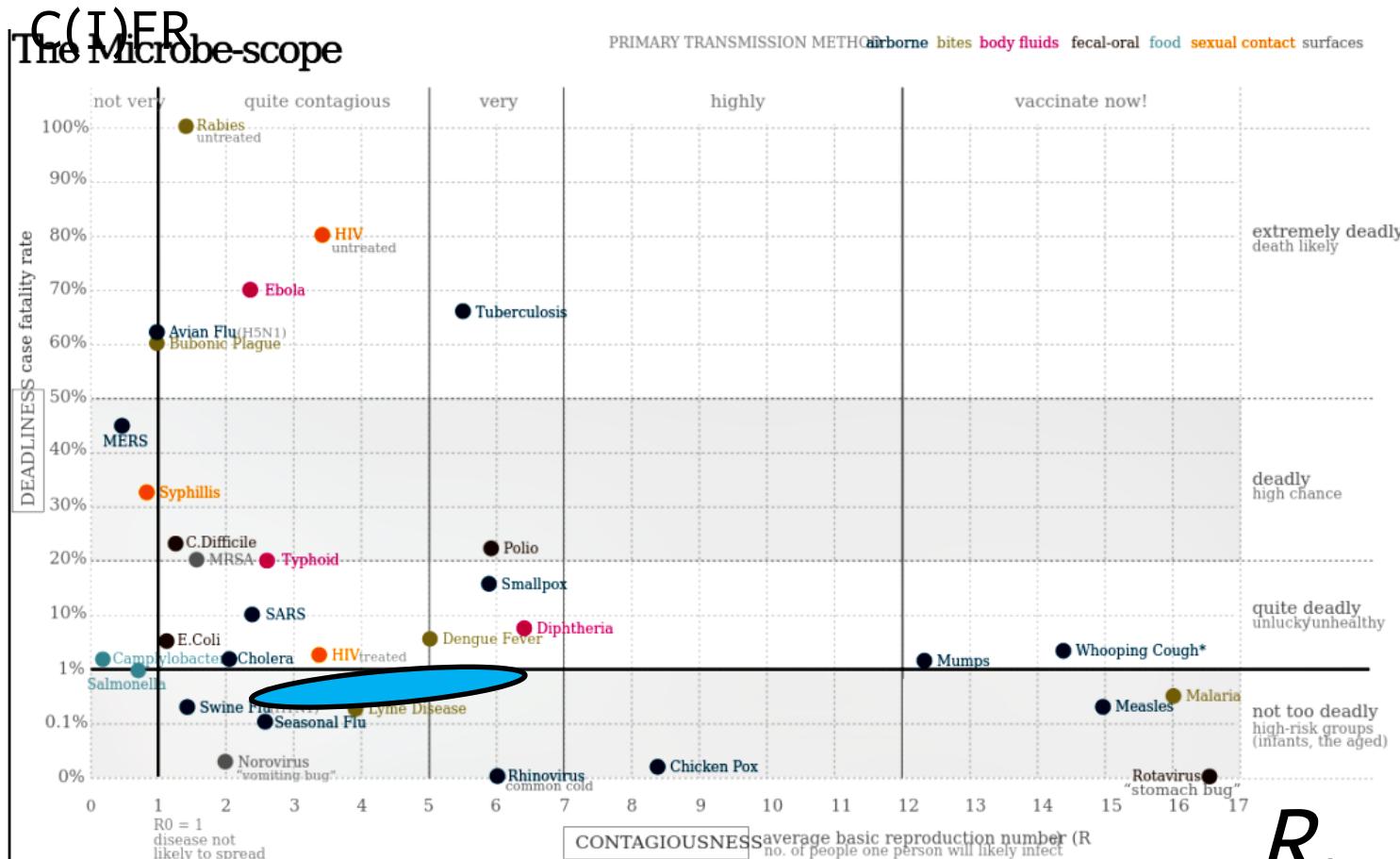
Figure 18. Mortalité toutes causes, tous âges confondus, de la semaine 36-2014 à la semaine 21-2021, France



Sources : Santé publique France, Insee

why did COVID-19 become pandemic

COVID-19



Fraser C et al. (2003) Factors that make an infectious disease outbreak controllable. *PNAS*

the natural history of COVID-19

historiques :

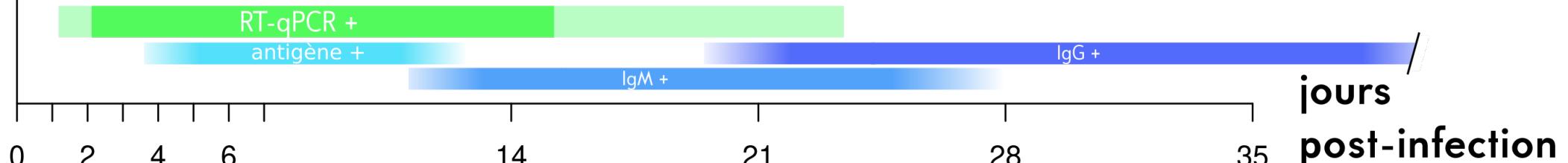
clinique



épidémiologique

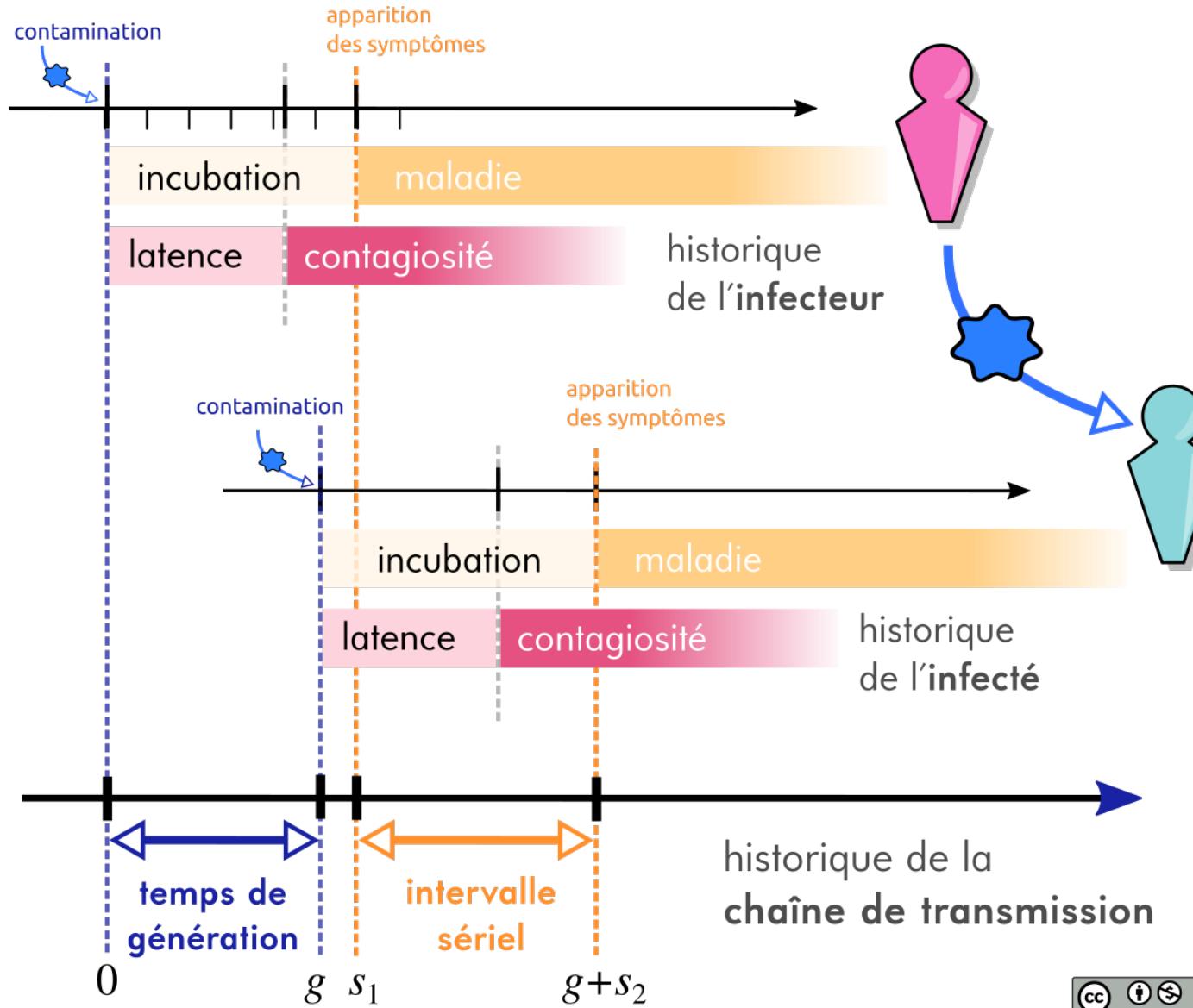


diagnostique



M.T.SOFONEA

generation time and serial interval



computing the pace of the epidemic

β_τ is the number of contaminations caused an index case on her/his τ^{th} day of infection

$\beta_1 + \beta_2 + \dots + \beta_\tau + \dots = \sum_{\tau > 0} \beta_\tau =: \mathcal{R}$ is the **reproduction number**

$w_\tau := \beta_\tau / \mathcal{R} \implies \sum_{\tau > 0} w_\tau = 1$ defines the **generation time distribution**

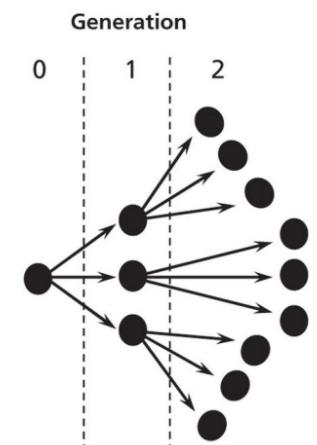
$y_t = \beta_1 y_{t-1} + \beta_2 y_{t-2} + \dots + \beta_\tau y_{t-\tau} + \dots = \mathcal{R} \sum_{\tau > 0} w_\tau y_{t-\tau}$ links incidence y_t on day t to those of previous days

$\widehat{\mathcal{R}}_t = y_t / \left(\sum_{\tau > 0} y_{t-\tau} w_\tau \right)$ is the estimated reproduction number at time t (as long as $y_t \propto y_t^{\text{true}}$)

An epidemic grows if and only if $\mathcal{R} > 1$ (**contagiosity**)

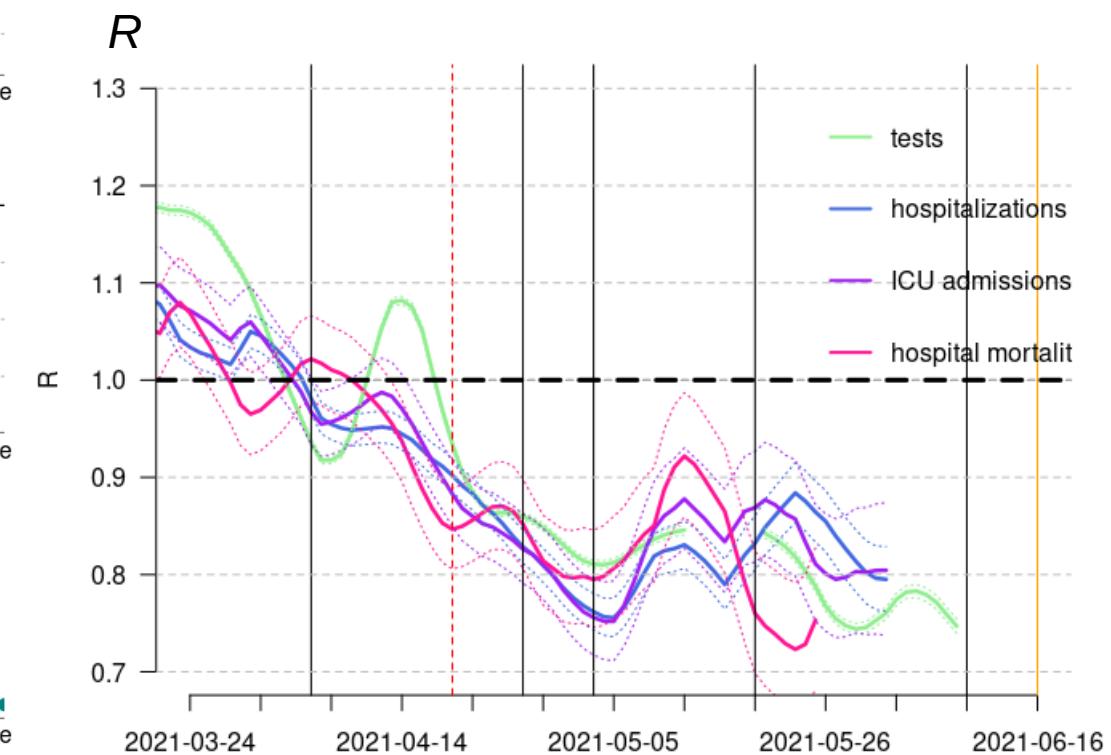
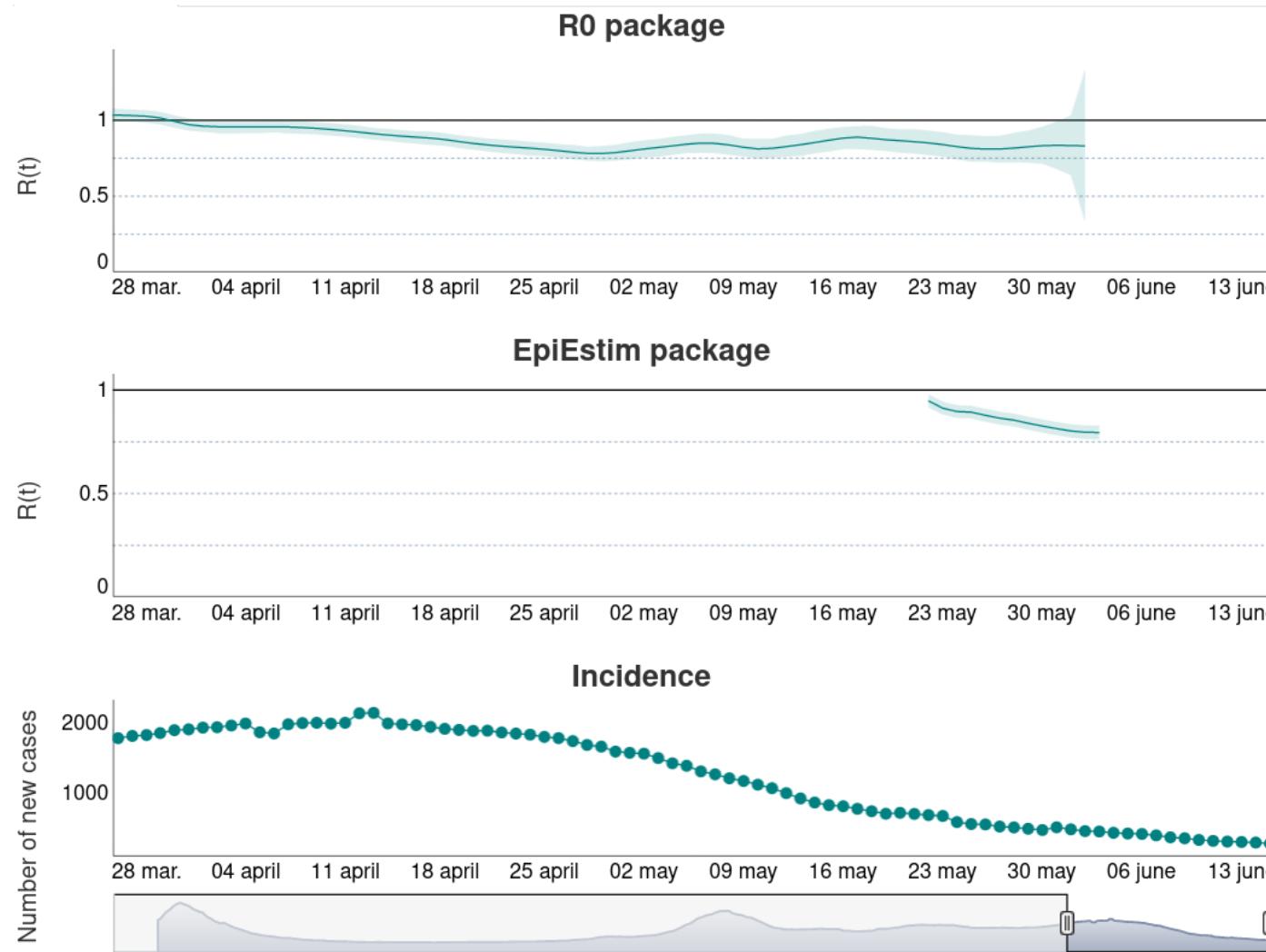
Assuming asymptotic behaviour, $y_t = \rho^t y_0 \implies \mathcal{R} \sum_{\tau > 0} \rho^{-\tau} w_\tau = 1$ (EULER-LOTKA equation)

$\log 2 / \log \rho =: T_2$ is the doubling/halving time (**speed**)



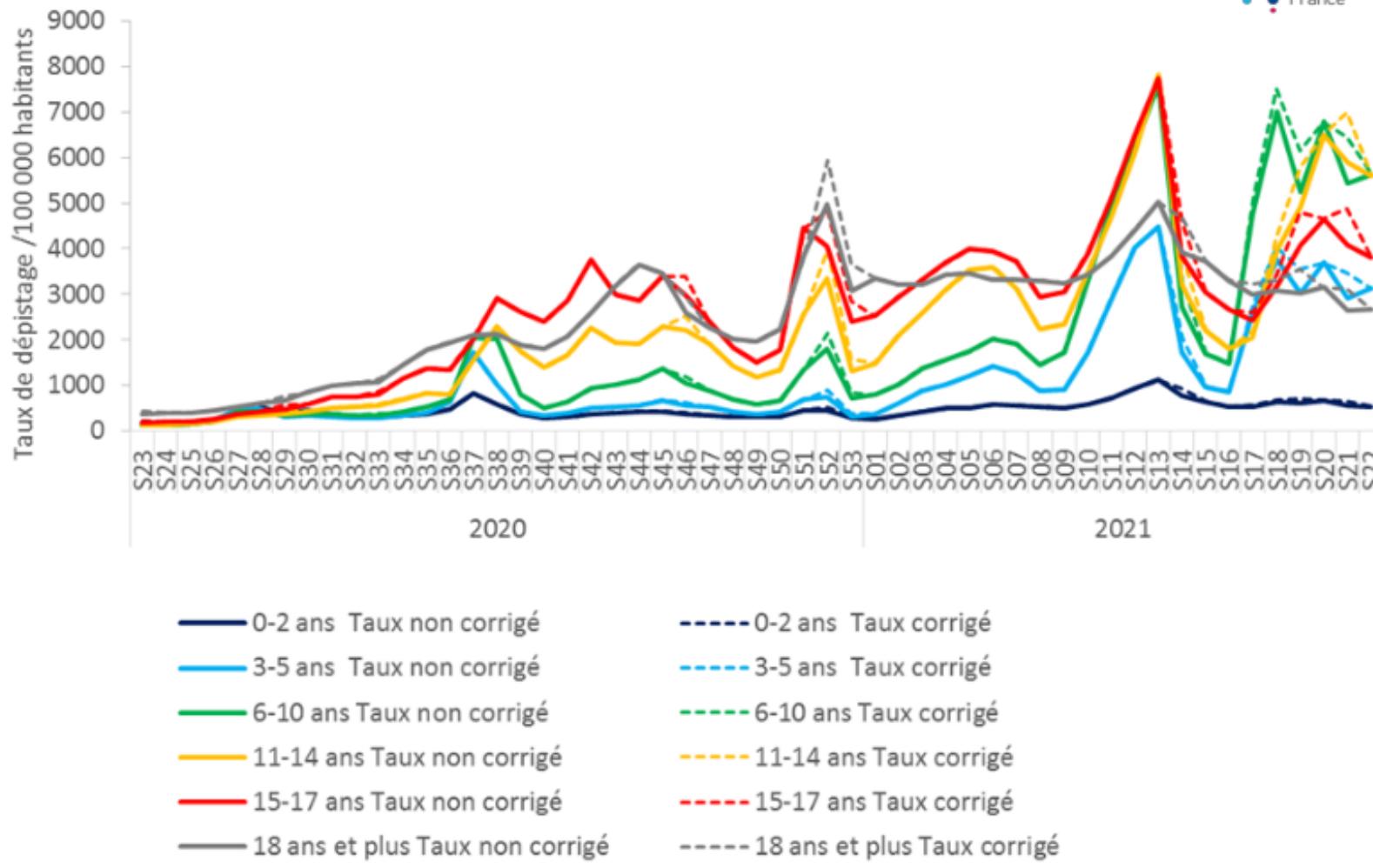
Initial phase of epidemic ($R_0 = 3$)
Pan-InforM 2009 (CMAJ)

estimating the reproduction number



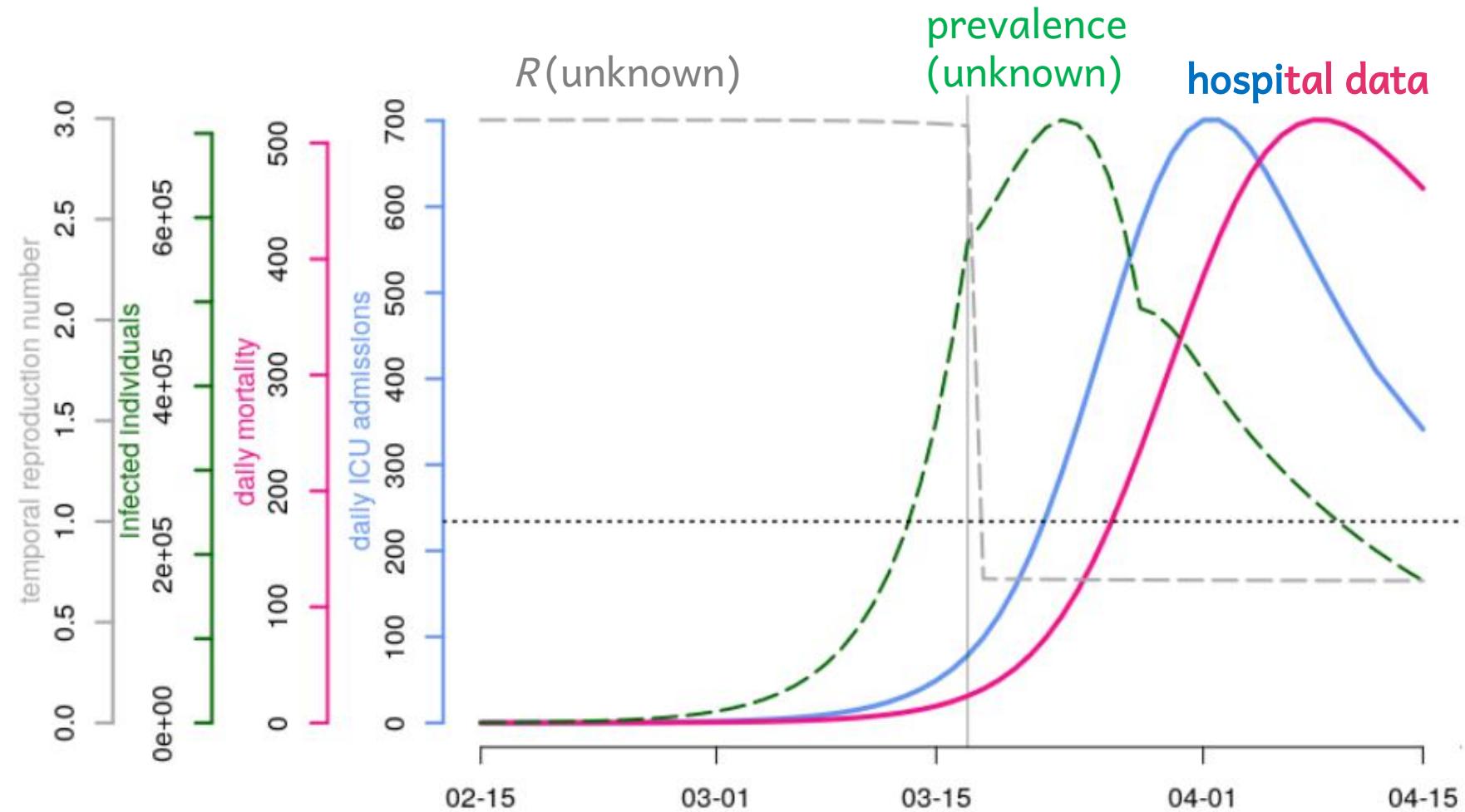
the non homogeneous detection issue

5b. Taux de dépistage



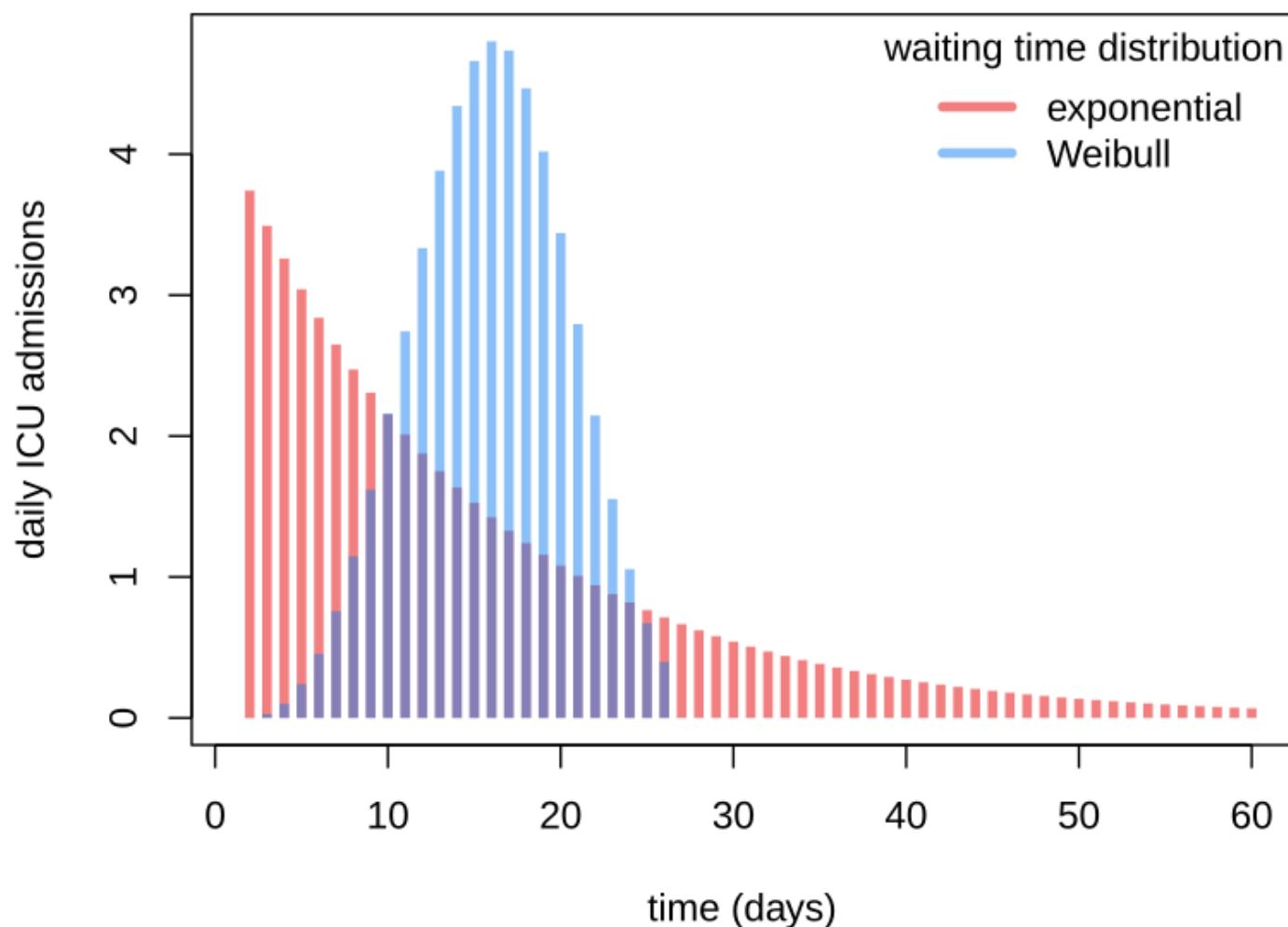
inferring the epidemic curve

- date and height of hospital capacity strain peak?
- lockdown effect?
- herd immunity?
- alternative restrictions? (age-specific, periodic, adaptative)

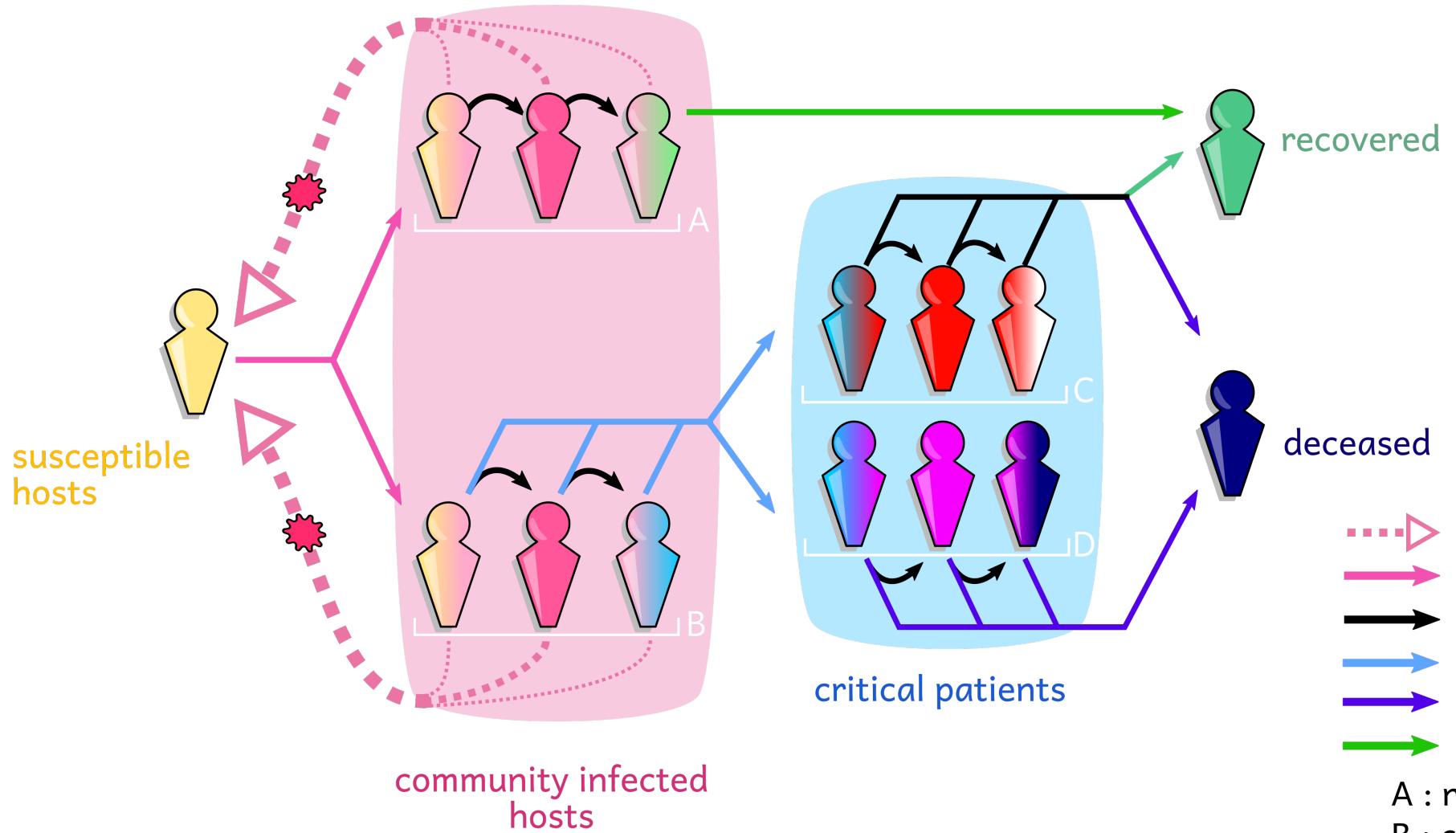


non-Markovian dynamics

ICU inflow from a 100 critical case cohort



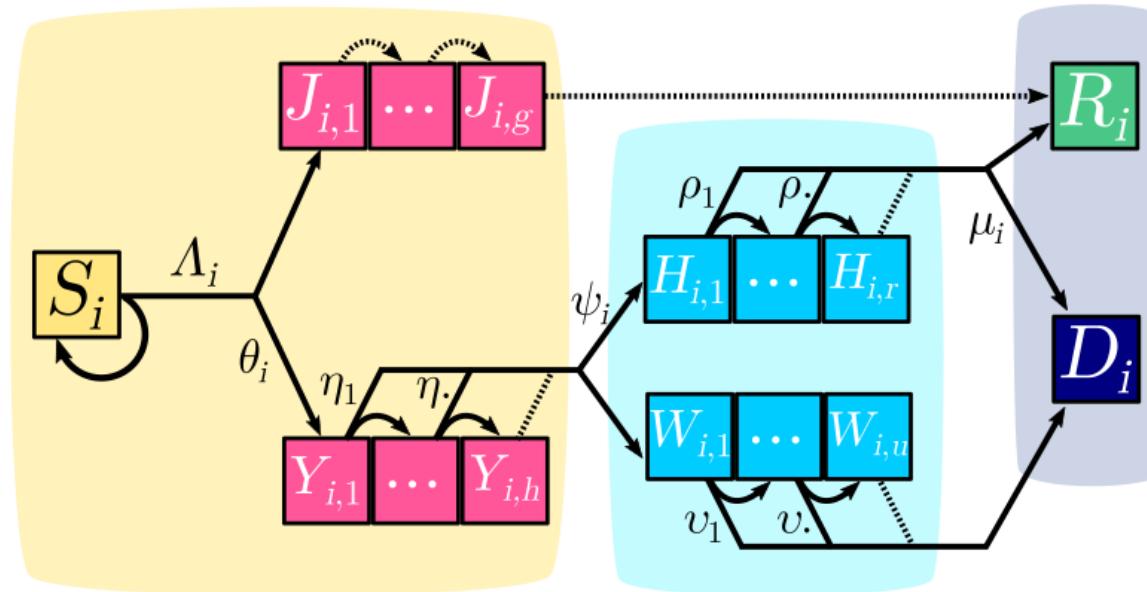
discrete time modelling



A : non critical cases
B : critical cases
C : in-ICU
D : out-ICU

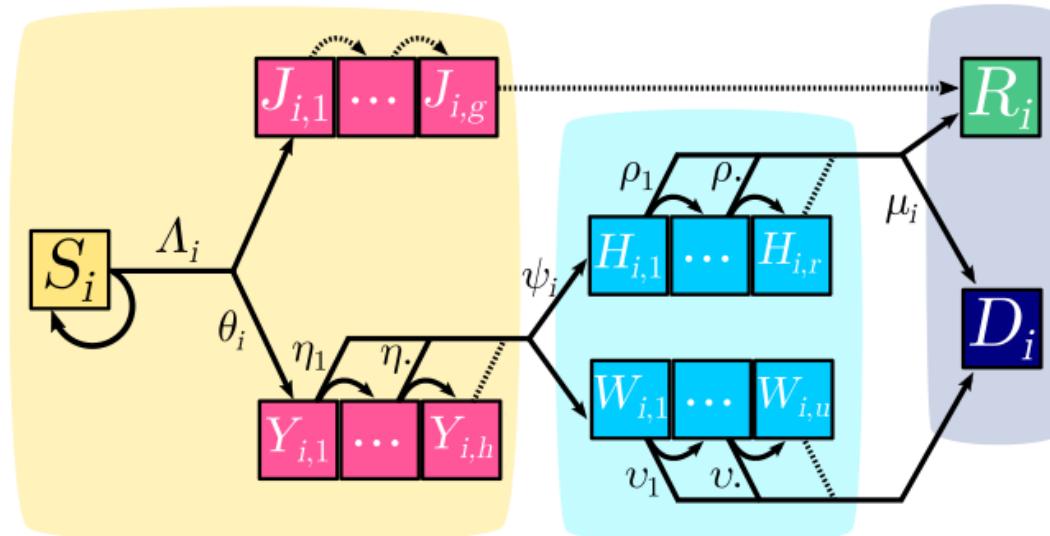
model structure

$S \equiv$ susceptible, $J \equiv$ non critical infectious, $Y \equiv$ critical inf.
 $H \equiv$ hospitalized critical patient (HCP) group 1,
 $W \equiv$ HCP2 (none or short ICU stay), $R \equiv$ recovered (immunized), $D \equiv$ dead



$i \equiv$ age group, $\Lambda \equiv$ force of infection, $\theta \equiv$ critical illness frequency
 $\eta \equiv$ contamination to hospitalization interval distribution
 $\psi \equiv$ long-stay ICU frequency among critical cases
 $\rho \equiv$ (HCP1) ICU length distribution, $\mu \equiv$ (HCP1) fatality rate
 $v \equiv$ (HCP2) hospitalization to death interval distribution

recurrence relation system



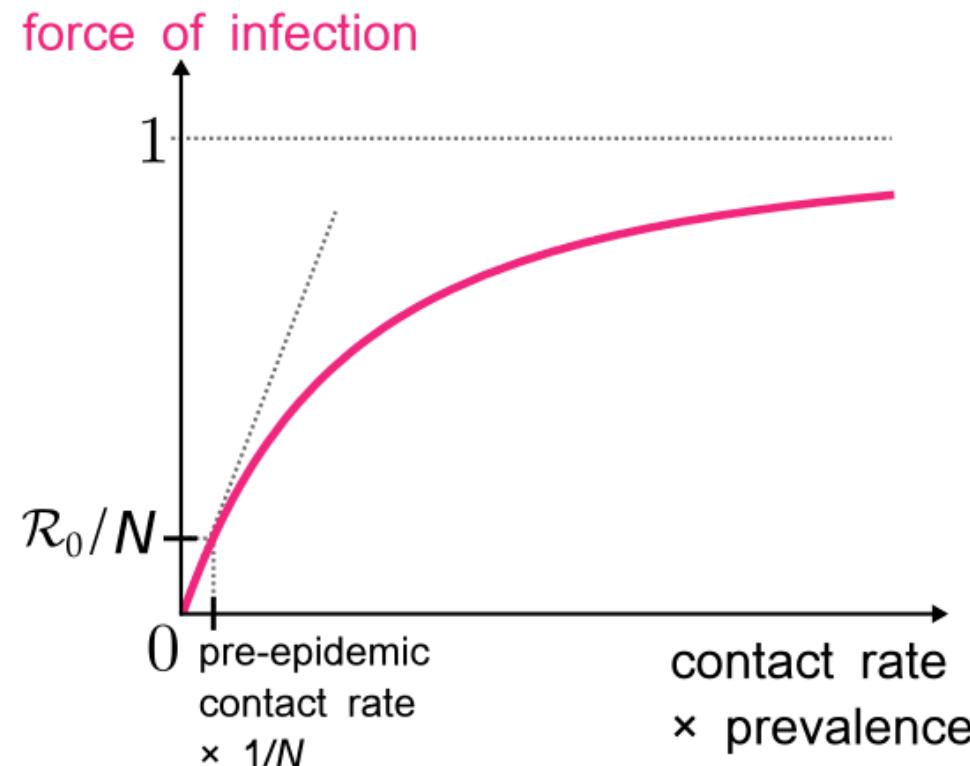
susceptible	$S'_i = (1 - \Lambda_i) S_i$		
non critical inf.	$J'_{i,1} = (1 - \theta_i) \Lambda_i S_i$	$J'_{i,k} = J_{i,k-1}$	$(1 < k < g)$
critical inf.	$Y'_{i,1} = \theta_i \Lambda_i S_i$	$Y'_{i,k} = (1 - \eta_{k-1}) Y_{i,k-1}$	$(1 < k < h)$
HCP1	$H'_{i,1} = \psi_i \sum_{k=1}^h \eta_k Y_{i,k}$	$H'_{i,k} = (1 - \rho_{k-1}) H_{i,k-1}$	$(1 < k < r)$
HCP2	$W'_{i,1} = (1 - \psi_i) \sum_{k=1}^h \eta_k Y_{i,k}$	$W'_{i,k} = (1 - v_{k-1}) W_{i,k-1}$	$(1 < k < u)$
recovered		$R'_i = R_i + J_{i,g} + (1 - \mu_i) \sum_{k=1}^r \rho_k H_{i,k}$	
dead		$D'_i = D_i + \sum_{k=1}^u v_k W_{i,k} + \mu_i \sum_{k=1}^r \rho_k H_{i,k}$	

force of infection

$\pi_{j,\tau} \equiv$ prevalence of j -individuals infected since τ days

$k_j \equiv$ per capita contact rate ; $\zeta_\tau \equiv$ discretized serial interval distribution

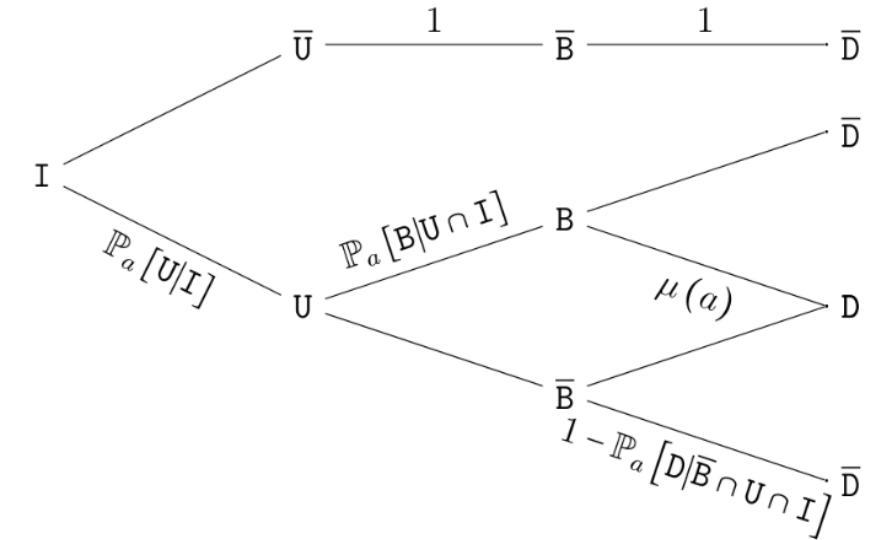
$$\Lambda_i(t) = 1 / \left(1 + 1 / \left(\mathcal{R}_0 \frac{k_i(t)}{k_i(0)} \sum_j \frac{k_j(t)}{k_j(0)} \sum_\tau \zeta_\tau \pi_{j,\tau}(t) \right) \right)$$



parametrisation

main input parameter	notation
basic reproduction number	\mathcal{R}_0
initiation day (YY-MM-DD)	t_0
lock-down control (%)	κ
critical case contamination to hospitalization interval expectation (days)	$\mathbb{E}[H]$
critical case contamination to hospitalization interval variance (days ²)	$\mathbb{V}[H]$
long ICU stay length expectation (days)	$\mathbb{E}[P]$
critical case hospitalization to death interval expectation (non long-stay ICU patients (days))	$\mathbb{E}[\Upsilon]$
infection fatality rate correction factor (%)	\mathfrak{C}_F
long-stay ICU fatality rate correction factor	\mathfrak{C}_M
long-stay ICU frequency correction factor	\mathfrak{C}_Ψ

$I \equiv$ being infected ; $U \equiv$ being hospitalized ; $B \equiv$ occupying an ICU bed for more than a day ; $D \equiv$ dying at the hospital from COVID-19.



$$\hat{\mu}(a) = \mathfrak{C}_M \mu(a) ; \hat{\psi}(a) = \mathfrak{C}_\Psi / \left(1 - \hat{\mu}(a) + \frac{\mathbb{P}_a[D|U \cap I]}{\mathbb{P}_a[B|U \cap I]} \right) ;$$

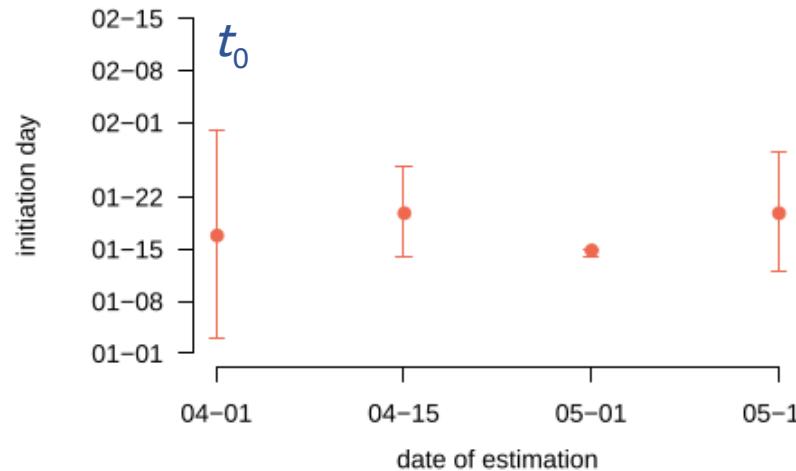
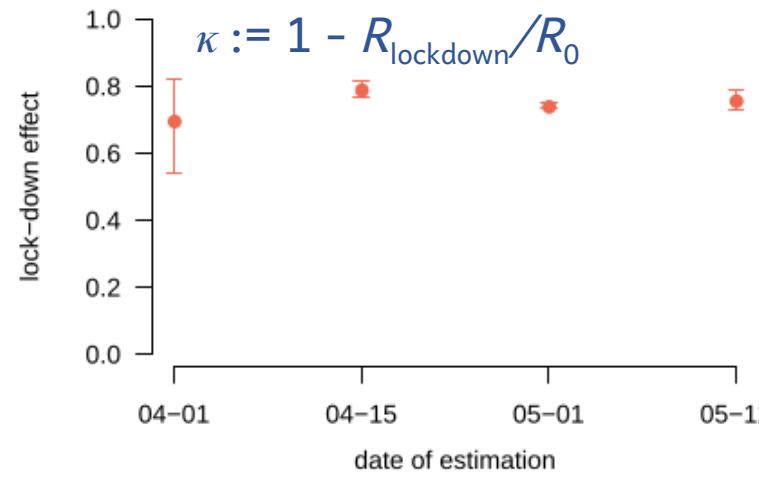
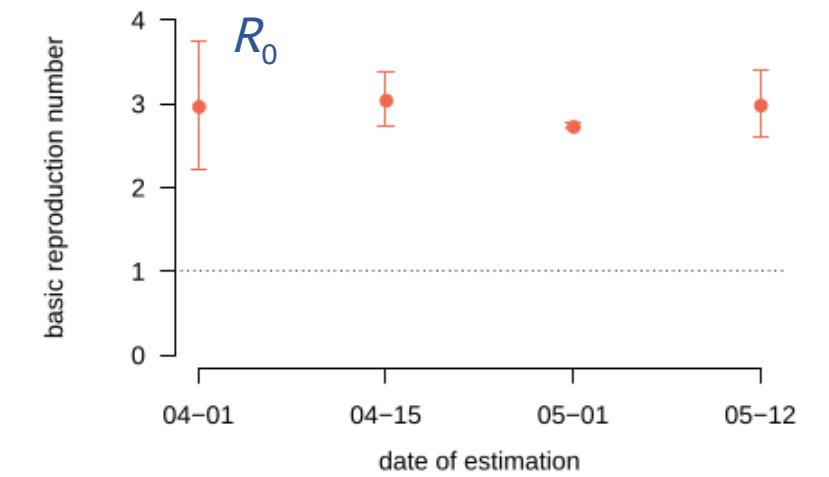
$$\hat{\theta}(a) = \frac{\mathfrak{C}_F \text{IFR}(a)}{1 - (1 - \hat{\mu}(a))\hat{\psi}(a)}.$$

fitting and early estimation of key parameters

Inference algorithm : R, mle2 (bbmle package).

Independent count process assumption : $X_{\text{obs}}(t) \sim \mathcal{P}(X_{\text{sim}}(t)) \rightarrow$ MLE and likelihood intervals.

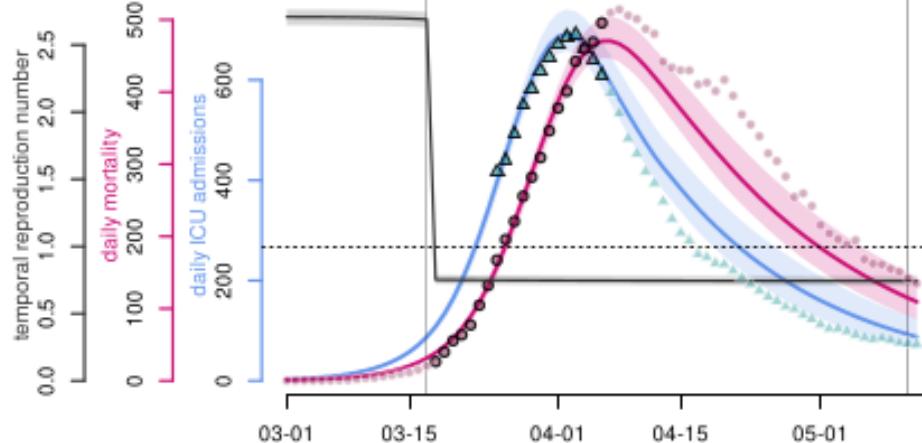
Random drawing around MLE $\xrightarrow{\text{Wilks}}$ 10^3 parameter sets \rightarrow display 2.5, 50 and 97.5% output quantiles.



MLE [95%-likelihood interval]
 $R_0 = 3.0 [2.6, 3.4] ; t_0 = \text{Jan 20} [12, 28] ; \kappa = 76 [73, 79] \%$

early estimation of key parameters

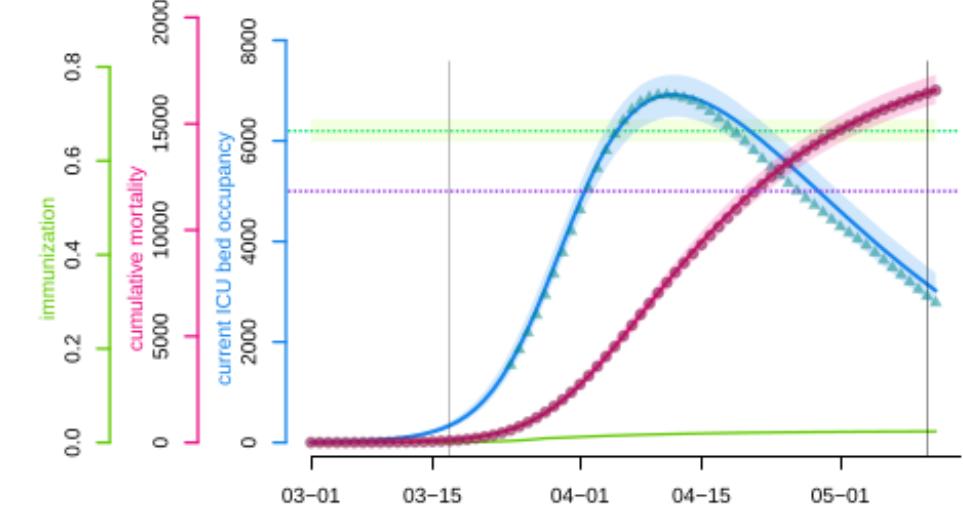
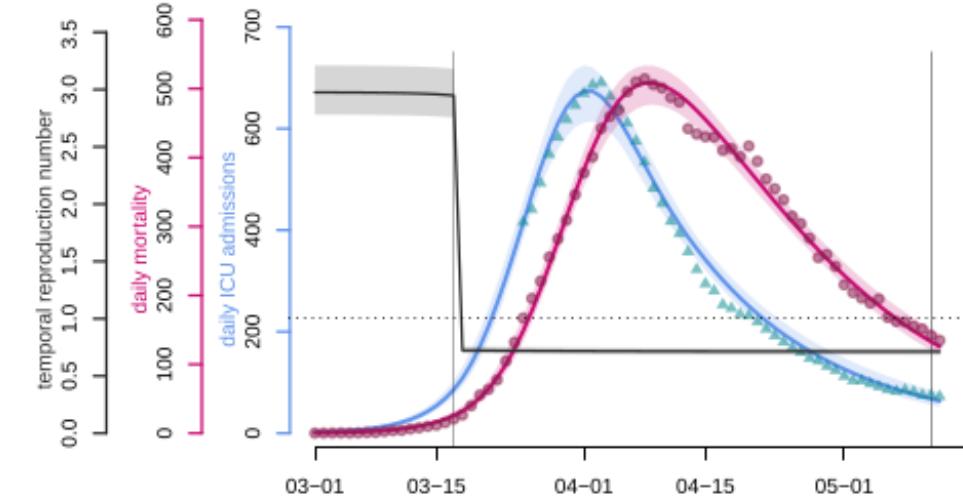
early forecasting
(data censored after Apr 7th 2020)



incidental

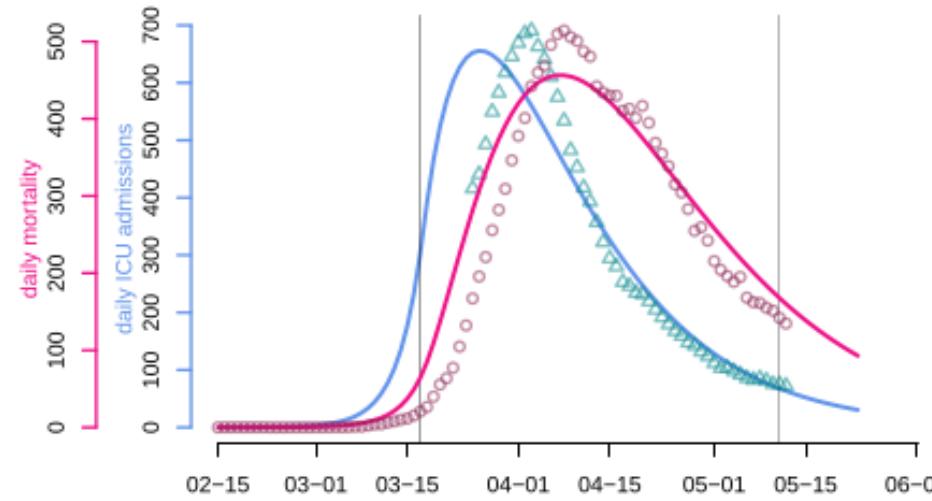
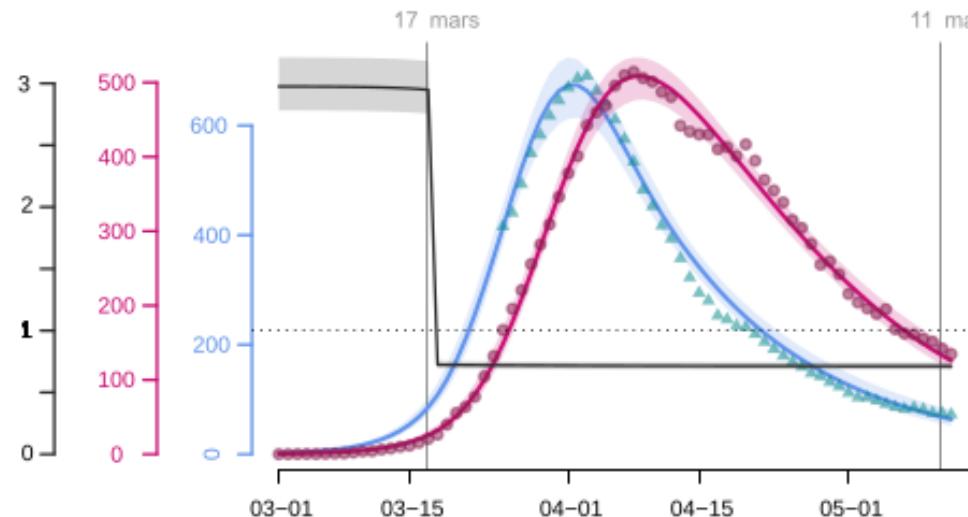
inertial

post-lockdown lifting (May 11th 2020)



comparison with the Markovian approach

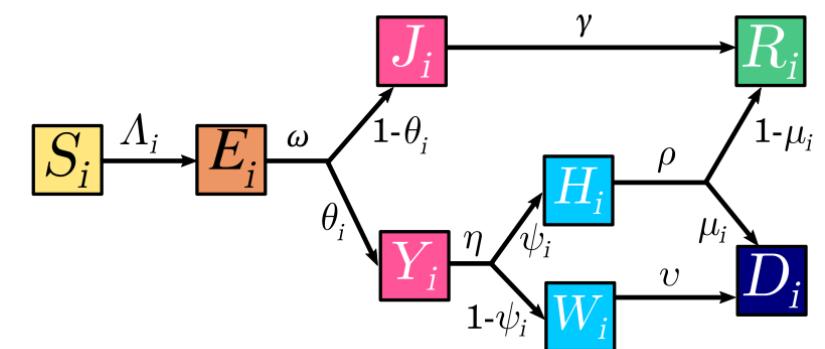
discrete time model vs SEAIR ODE



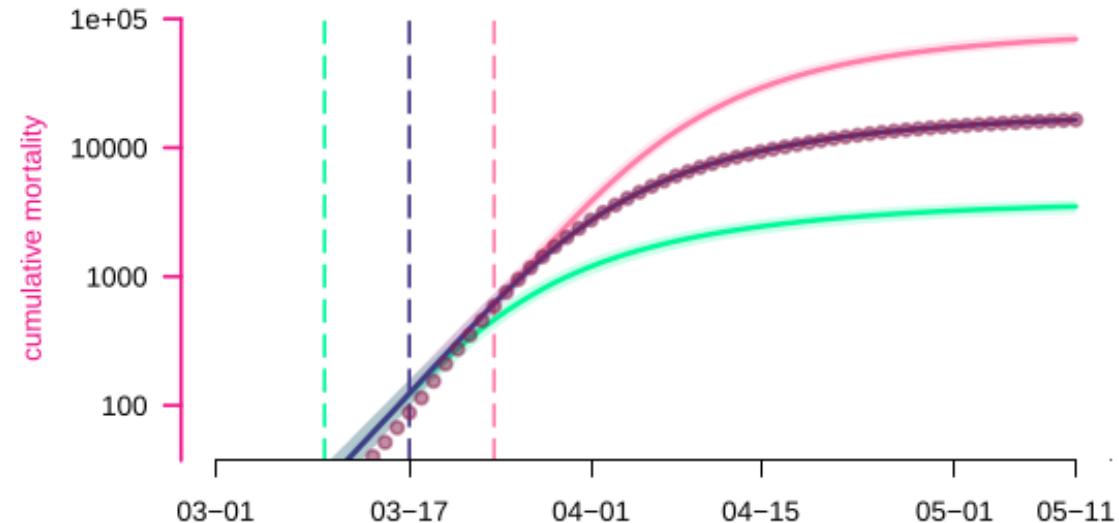
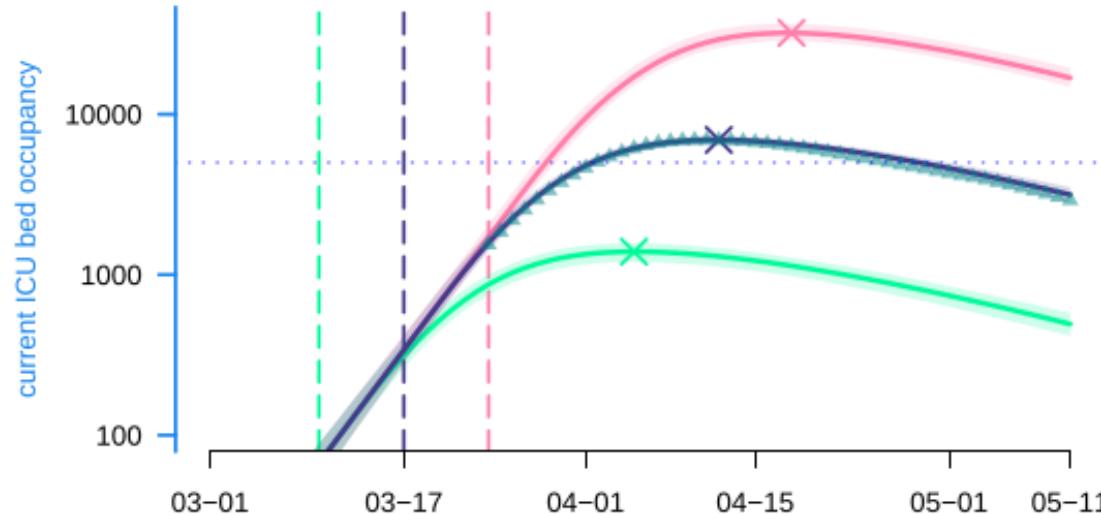
admissions en réanimation (nb/j)
médiane et IC_{0,95} des simulations
données lissées sur la semaine

décès hospitaliers (nb/j)
médiane et IC_{0,95} des simulations
données lissées sur la semaine

nombre de reproduction temporel
médiane et IC_{0,95} des simulations

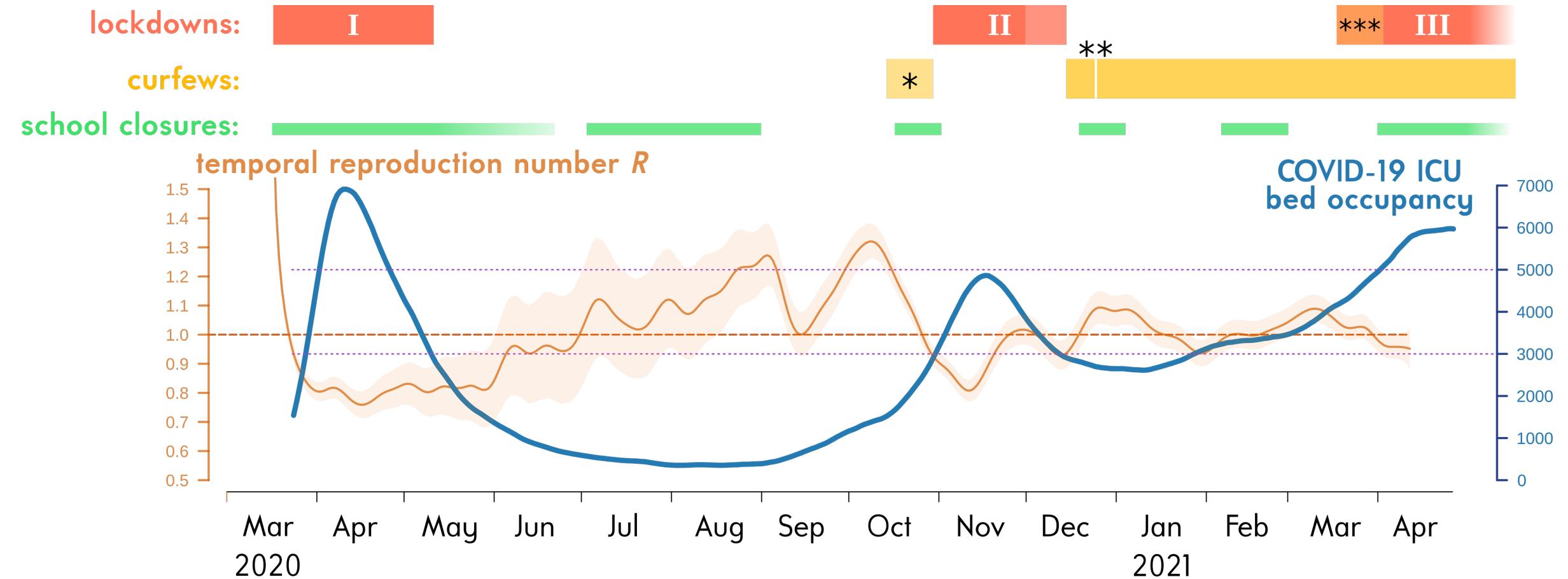


counterfactuals: lock-down timing



one week before: -13,300 [12,900 - 13,700] deaths,
one week after: +52,800 [45,800 - 61,500] deaths.

two waves and a high tide



per capita contact ratio

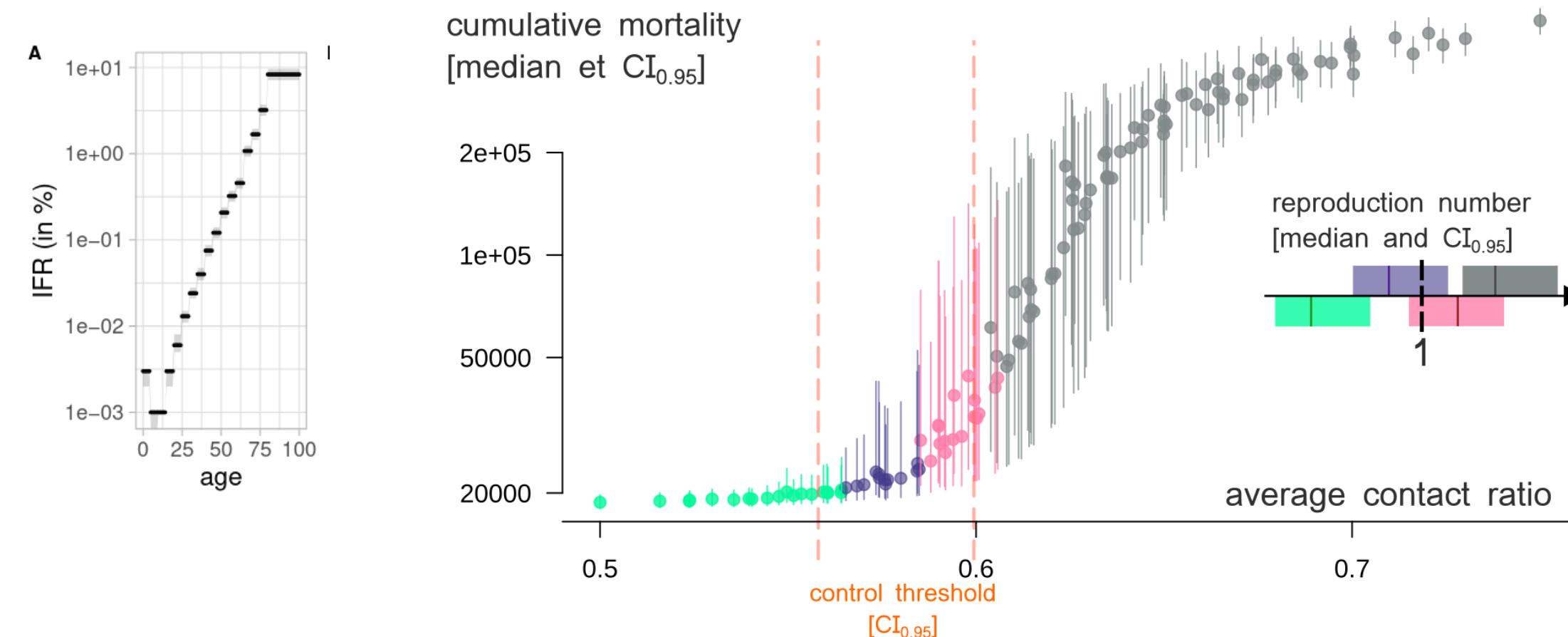
PCCR $c_i(t) := k_i(t)/k_i(0)$, average potentially infectious contacts an individual has each day, relative to the pre-epidemic baseline.

- proportionate mixing \Rightarrow encounter rate at t of an i - and a j -individuals $\propto c_i(t)c_j(t)$,
- $\overline{c_{\text{pre-lock}}} = 1$,
- $\overline{c_{\text{lock}}} = \sqrt{1 - \kappa} = 50\% [48, 52]\%$,
- threshold PCCR value corresponding to $\mathcal{R} = 1$ is $c_t = \mathcal{R}_0^{-\frac{1}{2}} = 58\% [56, 60]\%$.

$$\overline{c_{\text{lock}}} < c_{\text{under ctrl}} < c_t < c_{\text{out of ctrl}} < 1.$$

limited leverage of age-differential control

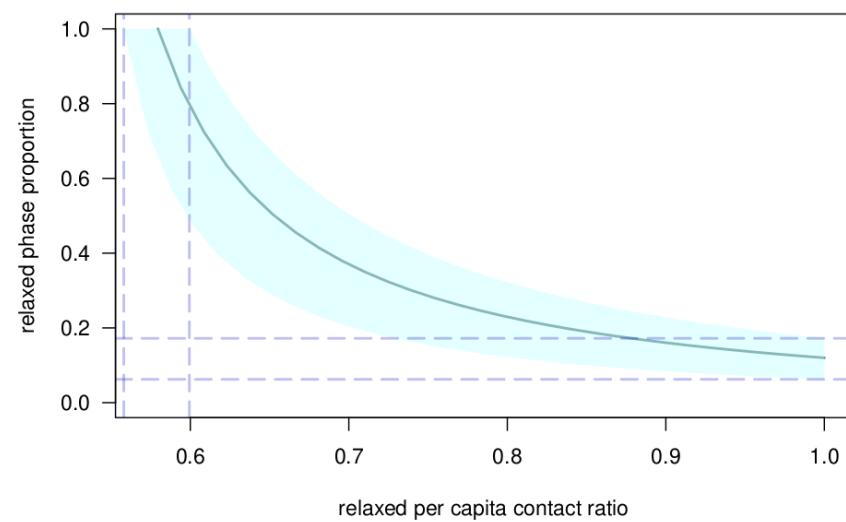
- three age groups with cut-offs 25 and 65 years old
- all combinations of per capita contact ratio (PCCR) = 50, 58, 62, 65, 75%



the unexplored partial periodic lockdowns

$p_r \equiv$ time proportion of relaxed phase ; $c_r, c_h \equiv$ relaxed, restricted PCCR
goal: maximizing p_r and c_r while $\bar{\mathcal{R}} = (p_r c_r^2 + (1 - p_r) c_h^2) \mathcal{R}_0 \leq 1$.

$$p_{r,\max} = \frac{\mathcal{R}_0^{-1} + \kappa - 1}{c_r^2 + \kappa - 1}.$$



the intensity-duration trade-off

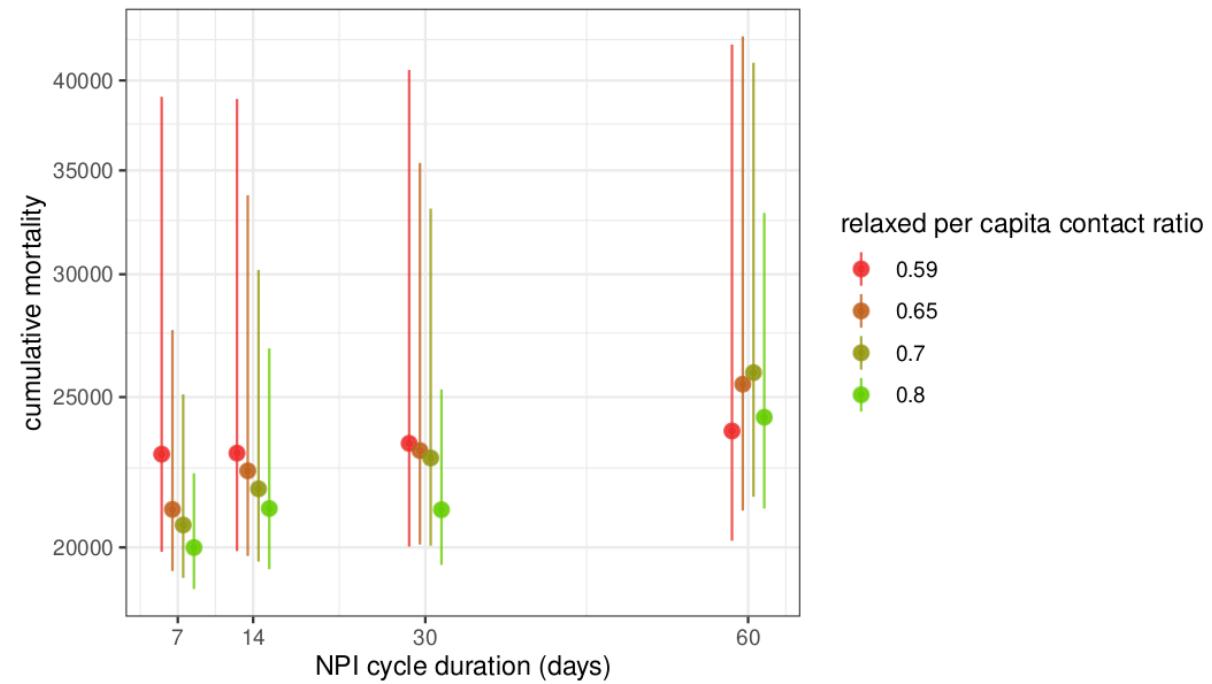
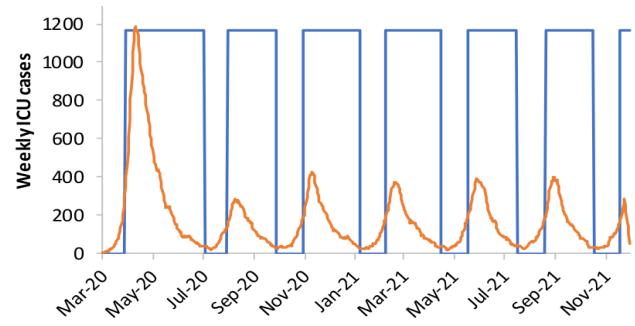


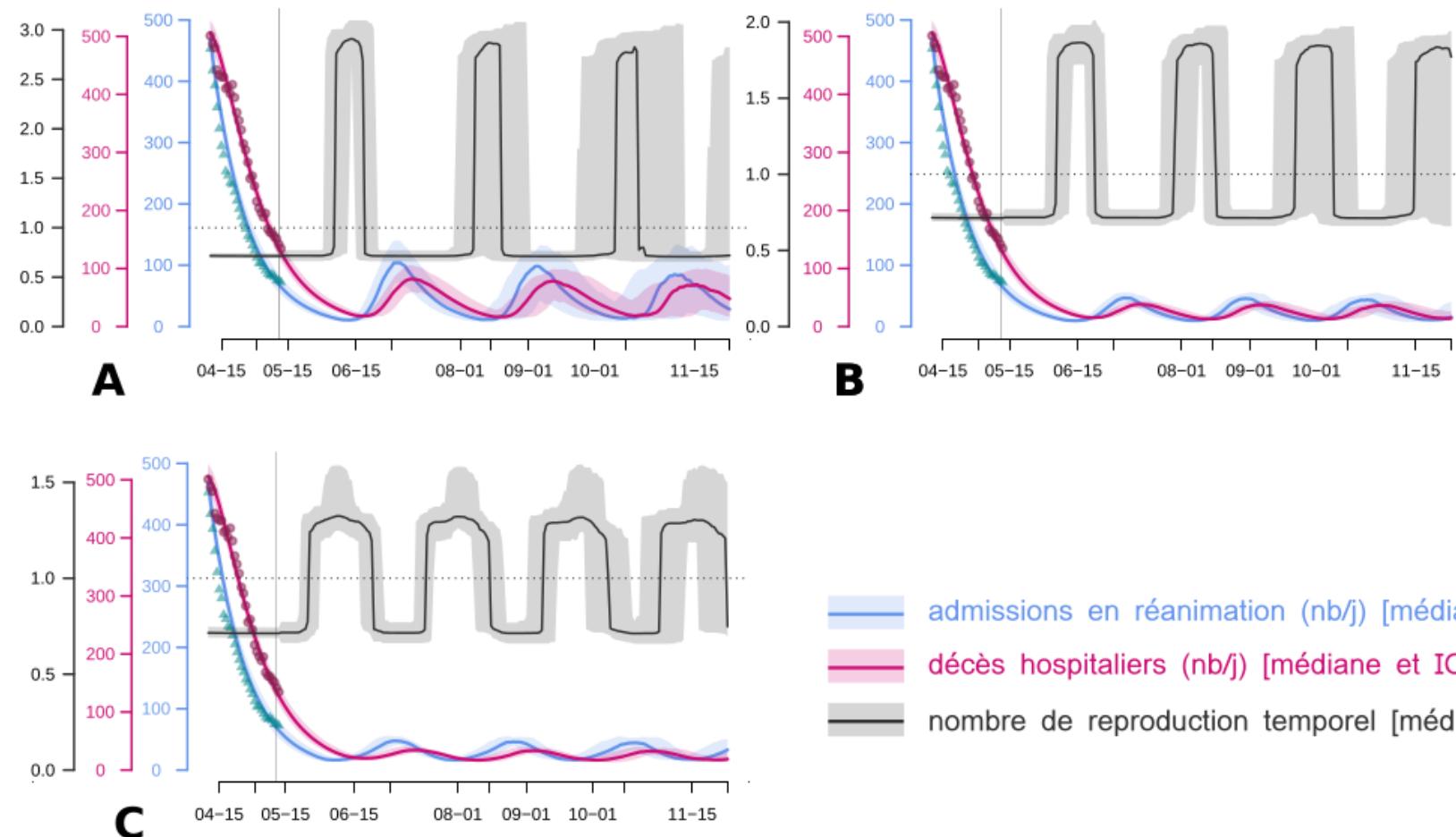
Figure 8: Effect of NPI cycle duration on cumulative mortality.

We show the cumulative mortality by the end of the year 2020 if a periodic control is implemented from May 12. For each of the four NPI cycle durations, the model is run for four values for c_r shown in different colors, with a proportion of time satisfying eq. 3 (truncated to an integer number of days). For example, in the red scenario, $p_{r,\max}(0.59) = 88\%$, so for the weekly cycle the relaxed phase was set to floor $(0.88 \cdot 7) = 6$ days.

adaptive control: illustrative cases



Ferguson et al. 20-03-16 'Report 9'



DIAT = daily ICU admission threshold (nationwide)

A: DIAT = 15, $c_r = 1$; **B:** DIAT = 15, $c_r = .8$,

C: DIAT = 5 (≥ 65 y.o.), 15 (< 25 y.o.), 30 (others), $c_r = .8$.

adaptive control: systematic analysis

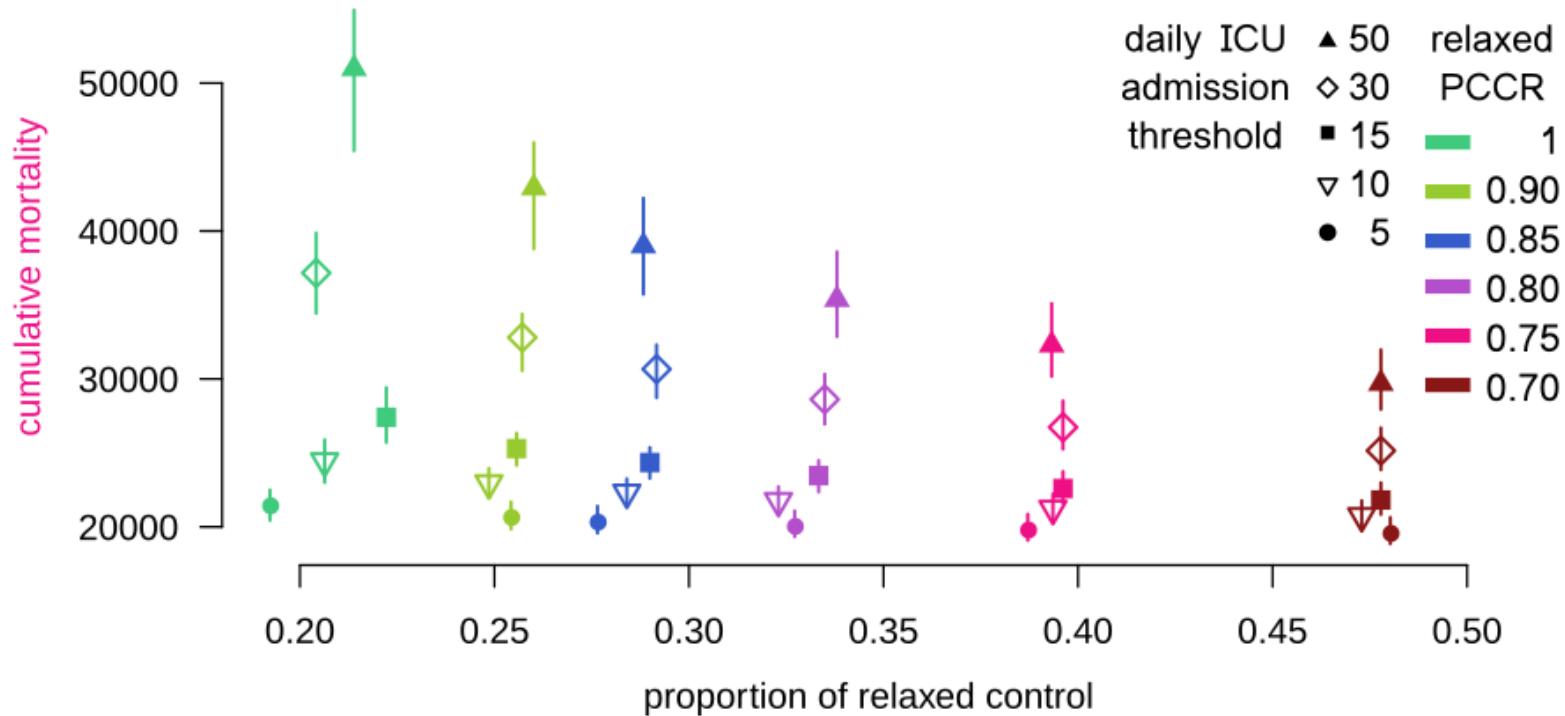
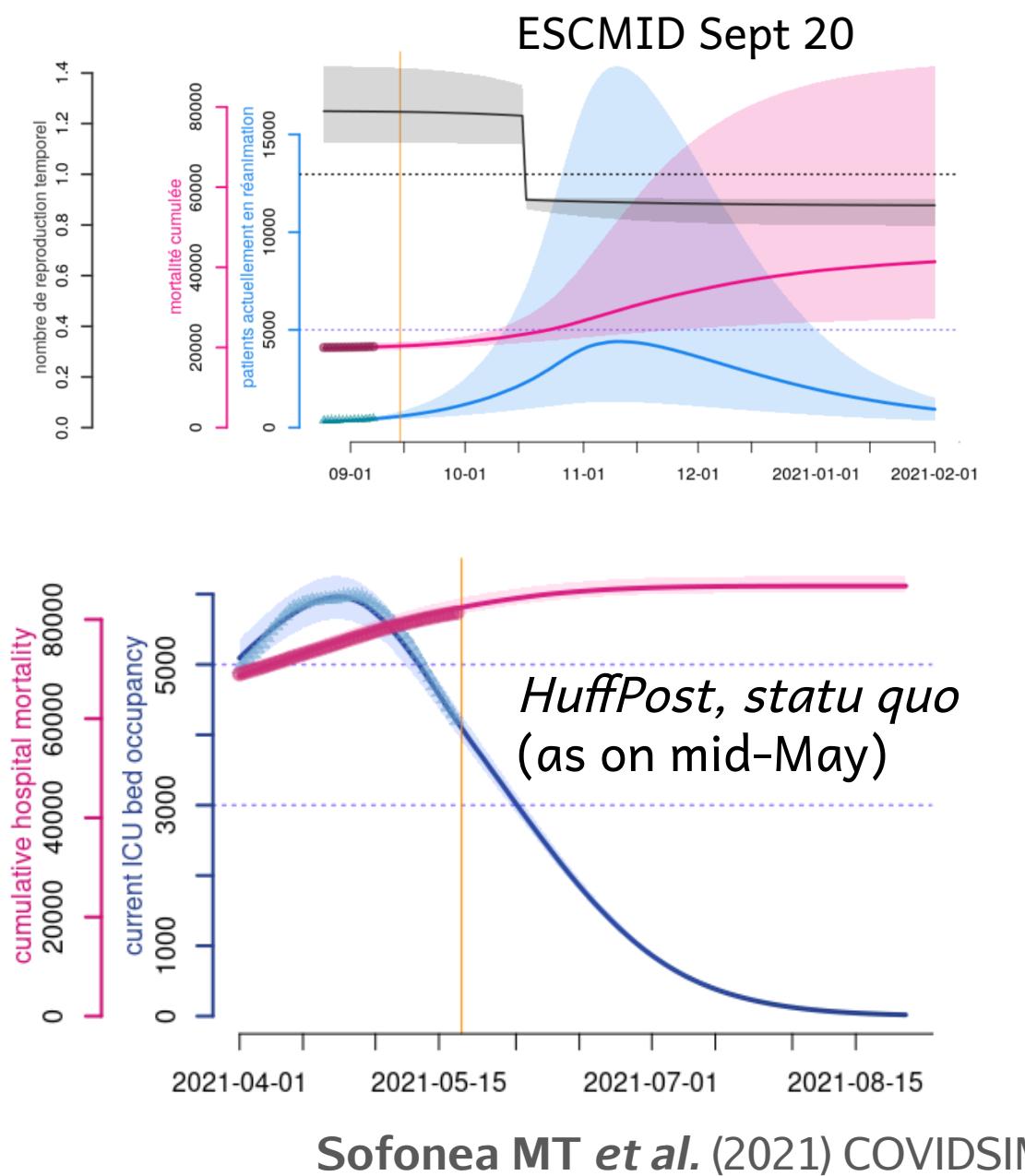
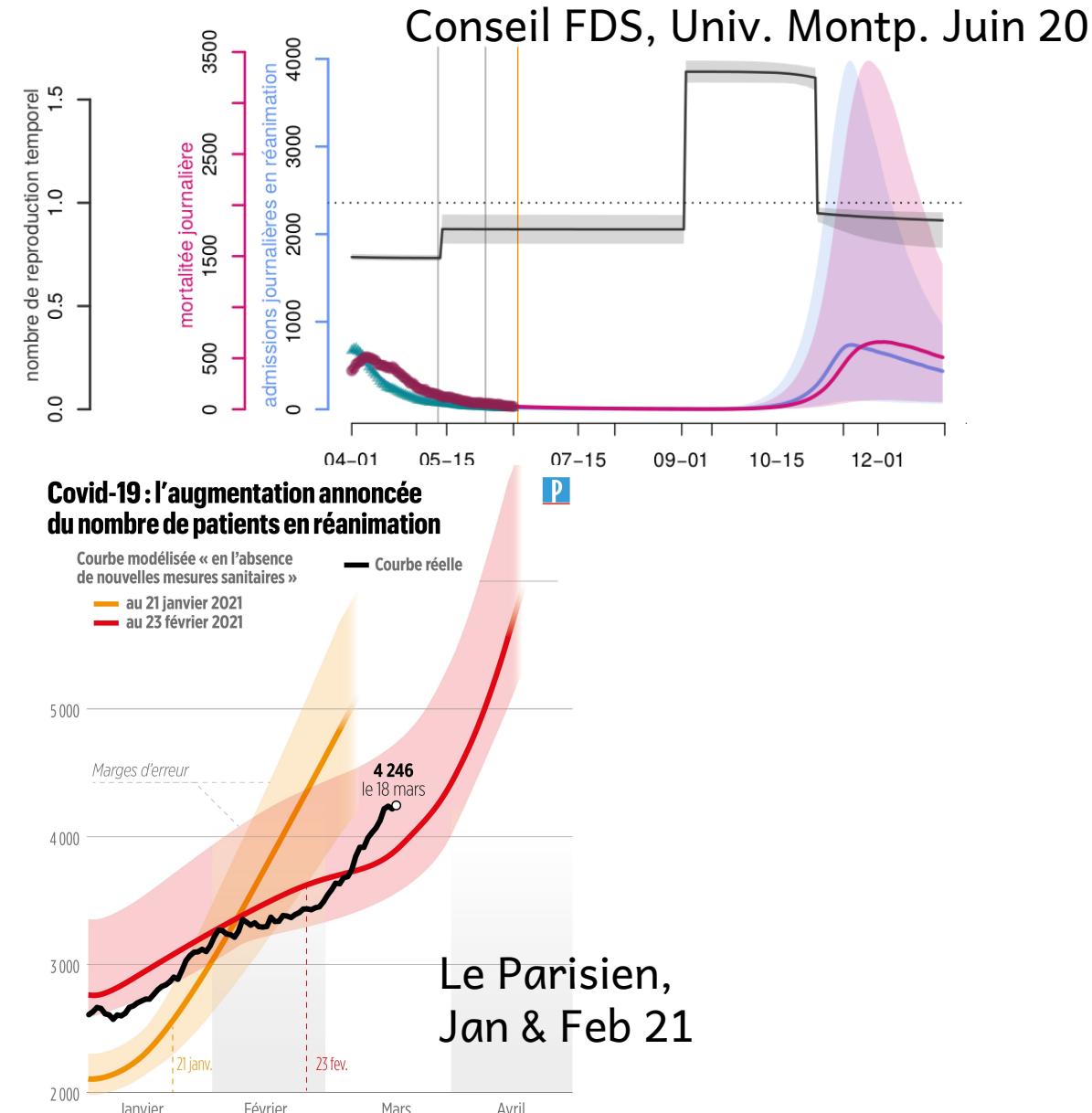


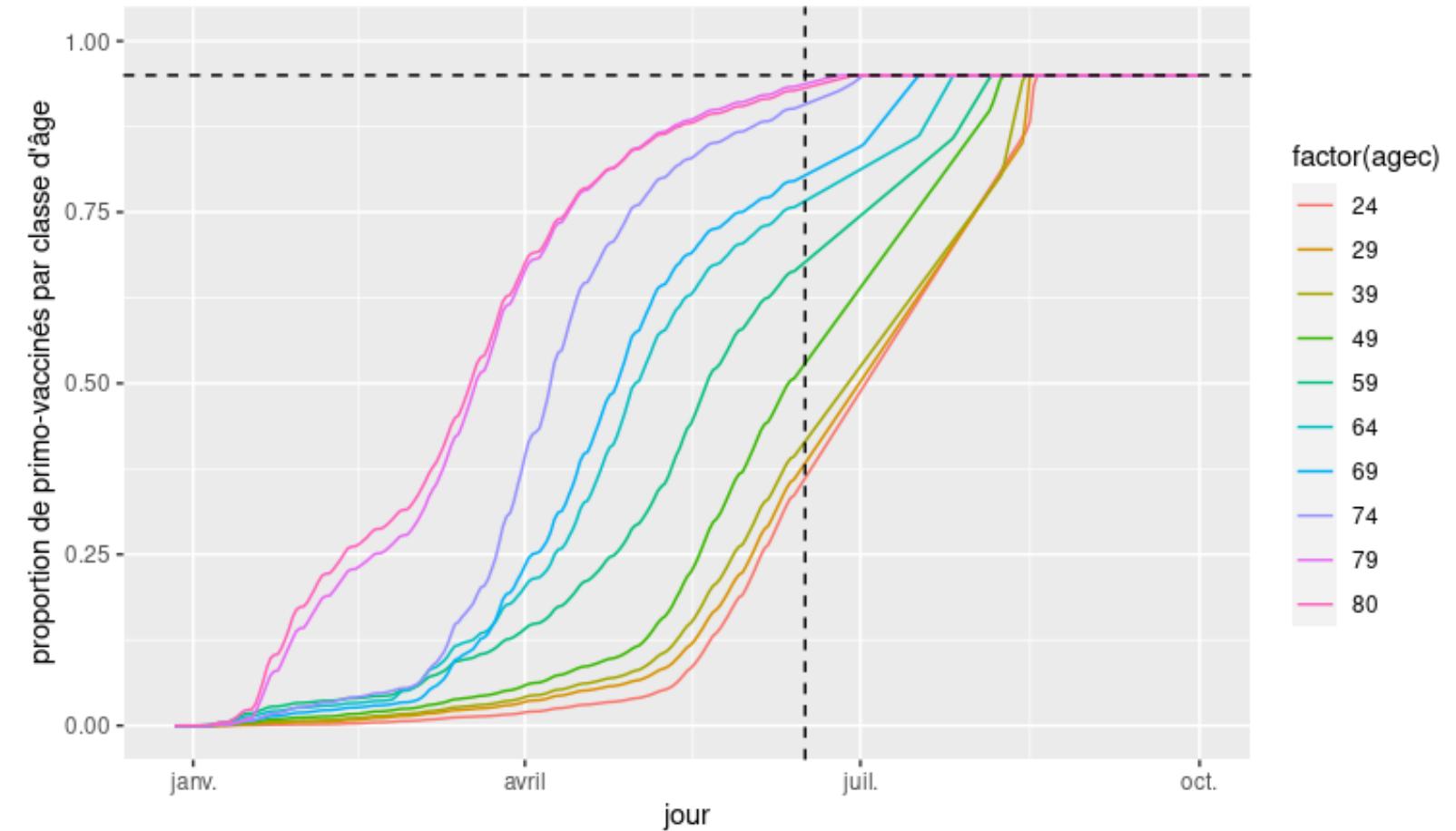
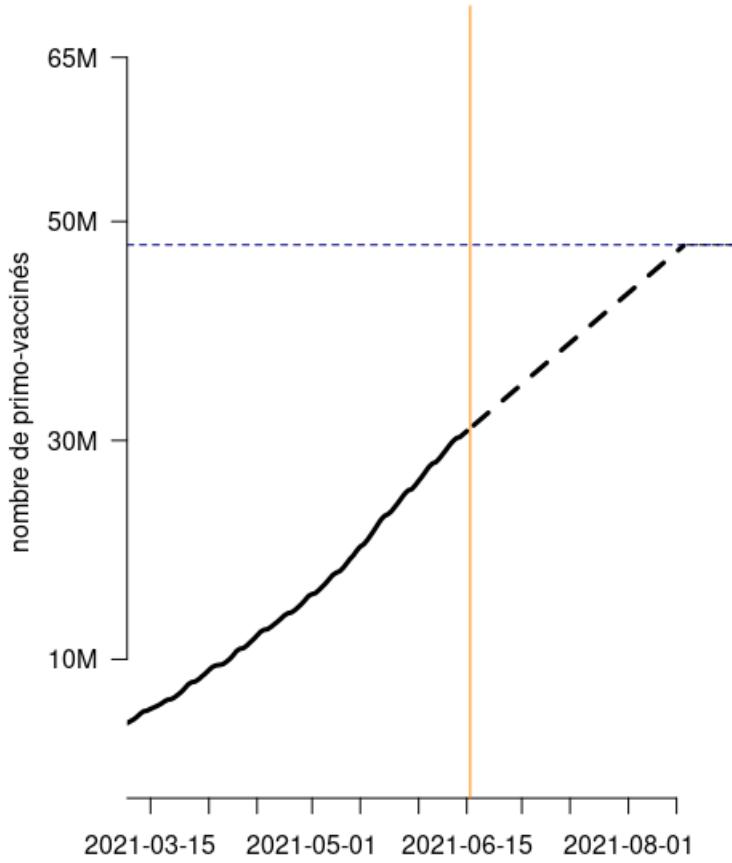
Fig. 7. Adaptive lock-down analysis

Each point represents a simulation of the model with adaptive lock-down implemented from May 12 given two input parameters: the lock-down triggering threshold in terms of daily ICU admissions nationwide and the PCCR during relaxed periods. The values of these parameters are respectively depicted by the shape and the color of the points, according to the legend. The abscissa show the median proportion of the time spent in the relaxed phase of each cycle. The ordinate represents the median final death toll by the end of the year, along with their 95%-confidence intervals (bars).

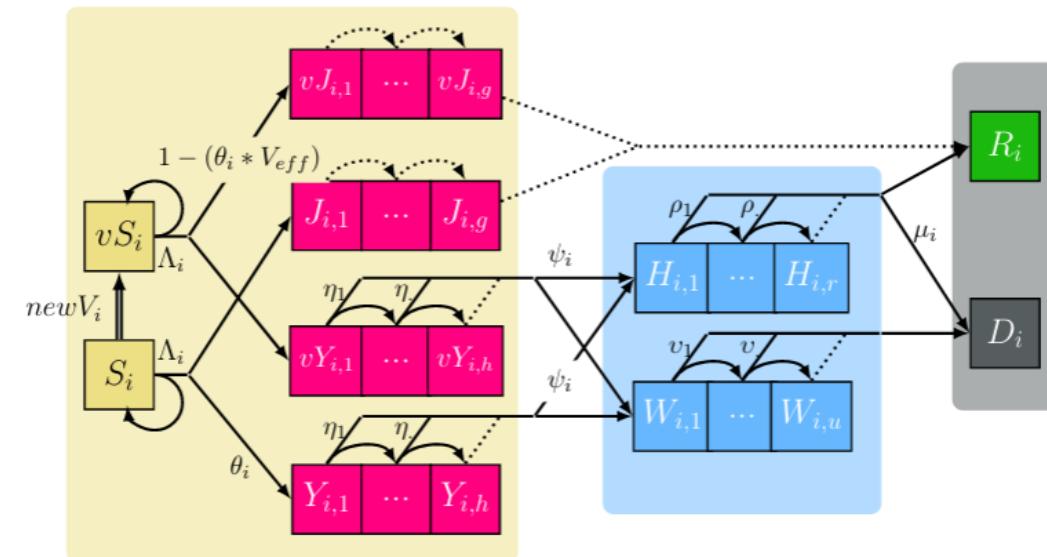
past projections



vaccine rollout



COVIDici



Veuillez choisir le niveau des données :

- Départementales
- Régionales
- Nationales

Tendance pour les projections :

- Courte
- Longue

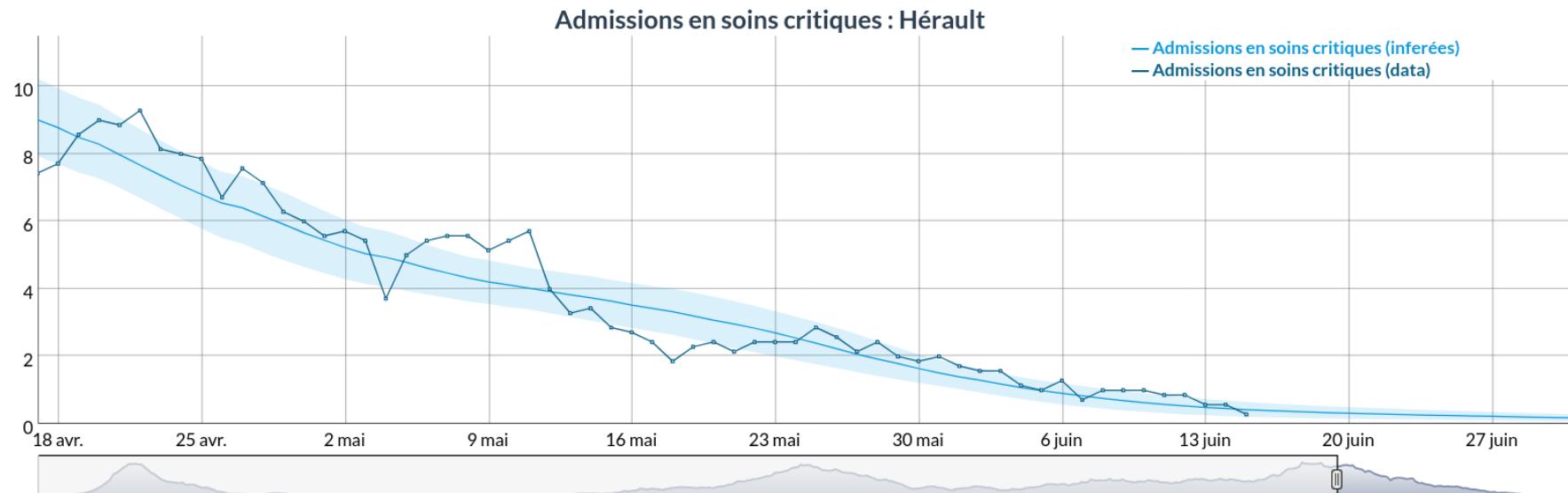
Afficher les dates des mesures restrictives nationales

Zone :

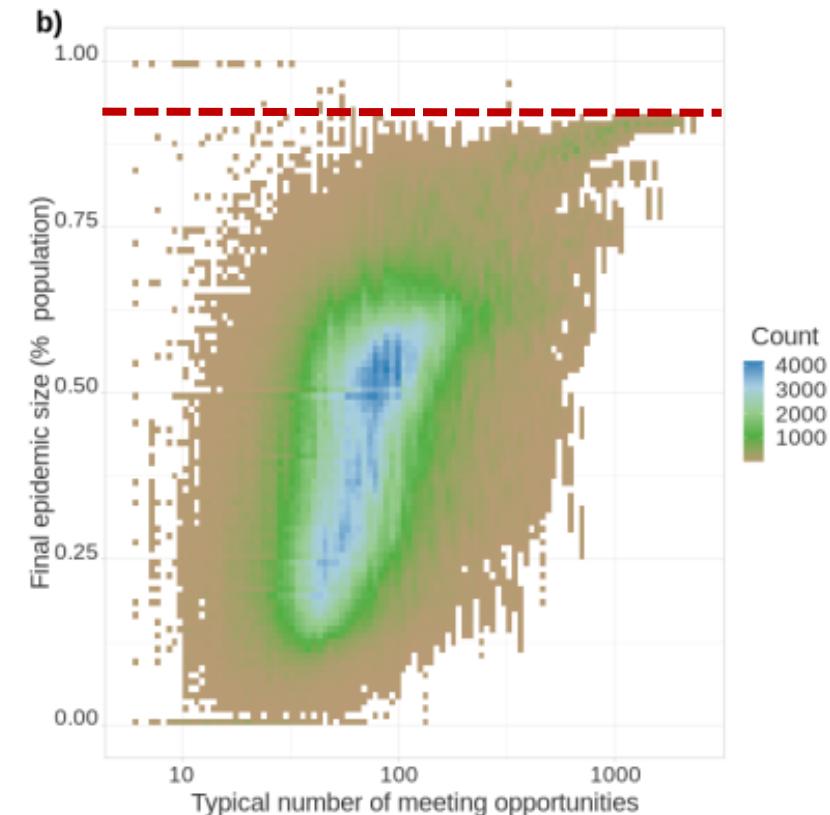
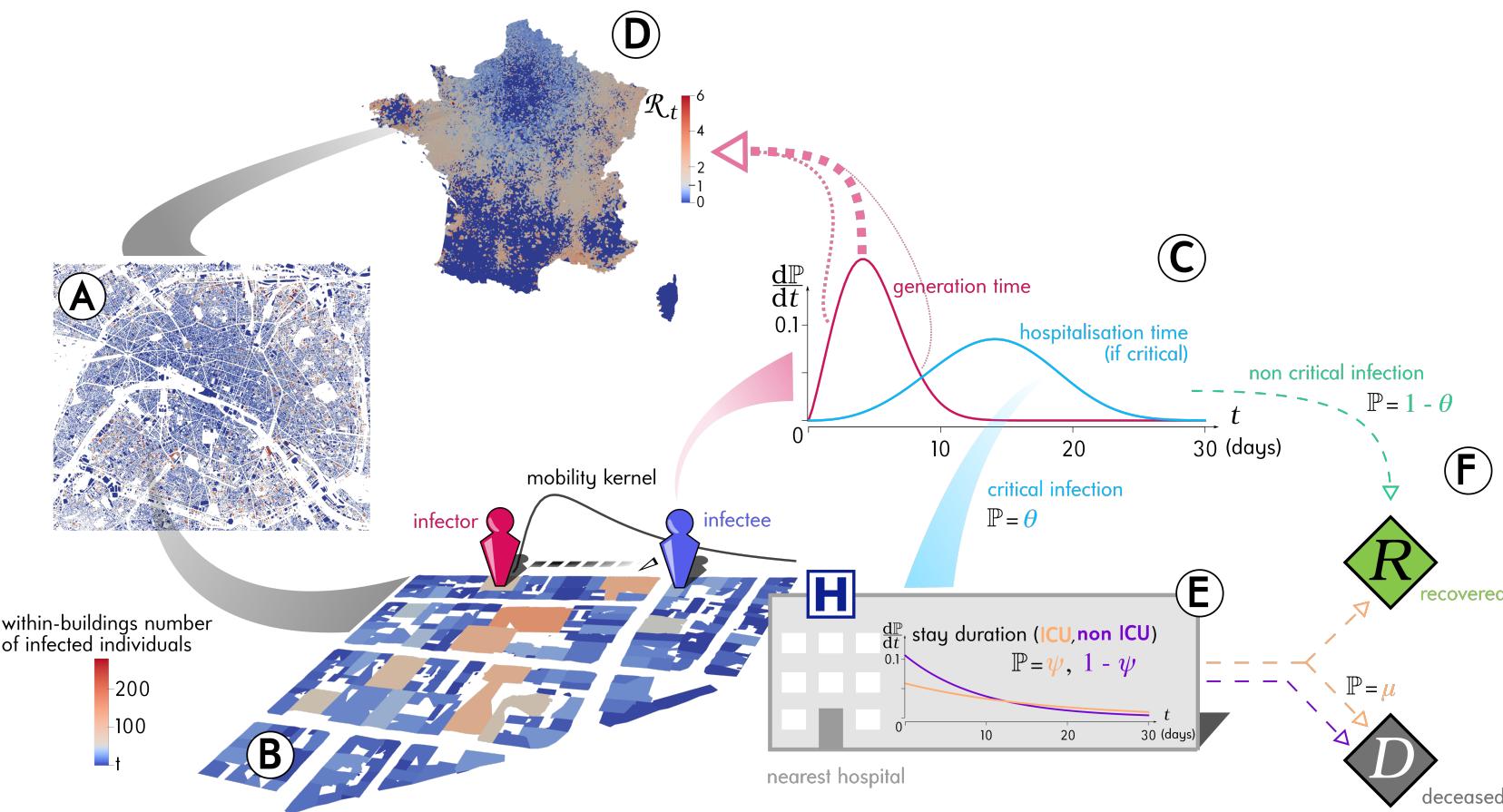
34 : Hérault

Veuillez choisir le type de données à visualiser :

- Admission en soins critiques



spatial considerations



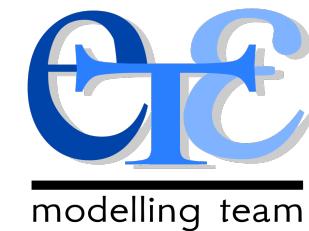
perspectives and acknowledgments

in brief

- discrete time is useful for combining memory effects, parsimony & computability,
- accounting for the age of infection is crucial for short-term prediction,
- age-related control variance shows limited leverage,
- importance of early strong responses and permanent moderate restrictions

future work

- time-varying contact matrices
- incidence-dependent NPIs
- implicit spatialisation



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@ete_fr

Samuel Alizon, Thomas Bénéteau,
Corentin Boennec, Marc Choisy,
Gonché Danesh, Ramsès Djidjou-Demasse
Baptiste Elie, Yannis Michalakis,
Bastien Reyné, Quentin Richard,
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Sonia Burrel

Olivier Thomine

Société Française
de Microbiologie

COVID-19 Study Group

parameter estimates

main input parameter	notation	maximum likelihood estimates	95% - likelihood interval
basic reproduction number	\mathcal{R}_0	2.99	[2.59, 3.39]
initiation day (YY-MM-DD)	t_0	20-01-20	[20-01-12, 20-01-28]
lock-down control (%)	κ	75.9	[72.9, 78.7]
critical case contamination to hospitalization interval expectation (days)	$\mathbb{E}[H]$	14.5	[13.6, 15.4]
critical case contamination to hospitalization interval variance (days ²)	$\mathbb{V}[H]$	20.0	[11.4, 30.9]
long ICU stay length expectation (days)	$\mathbb{E}[P]$	16.7	[14.9, 18.8]
critical case hospitalization to death interval expectation (non long-stay ICU patients (days)	$\mathbb{E}[\Upsilon]$	6.63	[6.19, 7.10]
infection fatality rate correction factor (%)	\mathfrak{C}_F	87.2	[85.8, 88.5]
long-stay ICU fatality rate correction factor	\mathfrak{C}_M	100.3	[100.2, 100.5]
long-stay ICU frequency correction factor	\mathfrak{C}_Ψ	93.8	[93.0, 94.5]

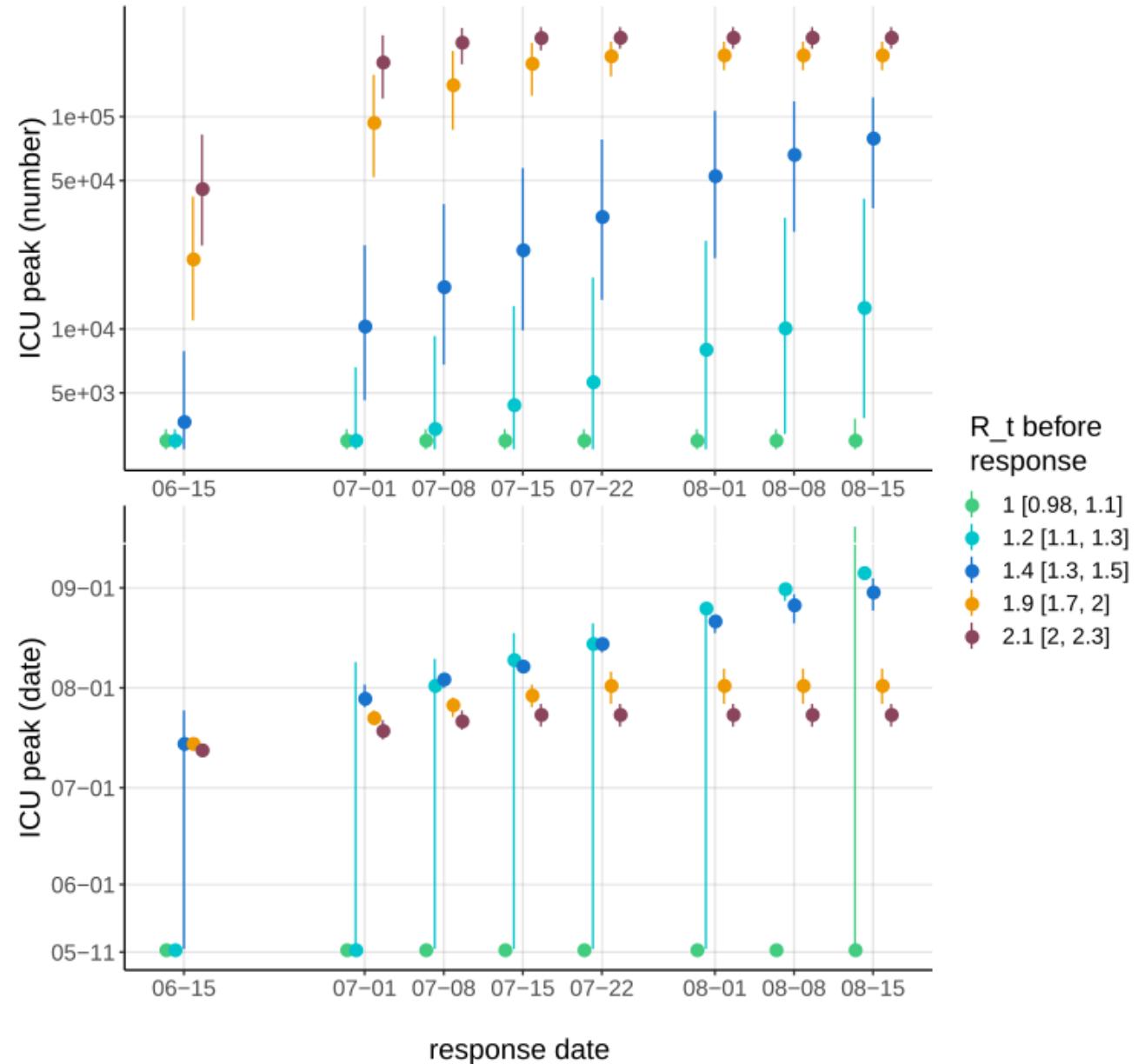
Table S-5: Maximum likelihood estimates and associated 95% - likelihood intervals for the ten input parameters. Details about the estimation procedure are provided in section [S2.7](#).

estimates with alternative IFR

main input parameter	notation	maximum likelihood estimates	95% - likelihood interval
basic reproduction number	\mathcal{R}_0	3.33	[3.16, 3.50]
initiation day (YY-MM-DD)	t_0	20-01-22	[20-01-20, 20-01-25]
lock-down control (%)	κ	77.6	[76.4, 78.7]
critical case contamination to hospitalization interval expectation (days)	$\mathbb{E}[H]$	14.0	[13.7, 14.3]
critical case contamination to hospitalization interval variance (days ²)	$\mathbb{V}[H]$	19.4	[18.2, 20.6]
long ICU stay length expectation (days)	$\mathbb{E}[P]$	19.2	[18.4, 19.9]
critical case hospitalization to death interval expectation (non long-stay ICU patients (days)	$\mathbb{E}[\Upsilon]$	9.43	[9.20, 9.66]
infection fatality rate correction factor (%)	\mathfrak{C}_F	62.0	[61.0, 63.0]
long-stay ICU fatality rate correction factor	\mathfrak{C}_M	110.0	[74.4, 146.0]
long-stay ICU frequency correction factor	\mathfrak{C}_Ψ	100.2	[100.1, 100.2]

Table S-6: Maximum likelihood estimates and associated 95% - likelihood intervals for the ten input parameters using alternative IFR values. Details about the estimation procedure are provided in section [S2.7](#).

post-lock-down response delay



adaptive control: threshold

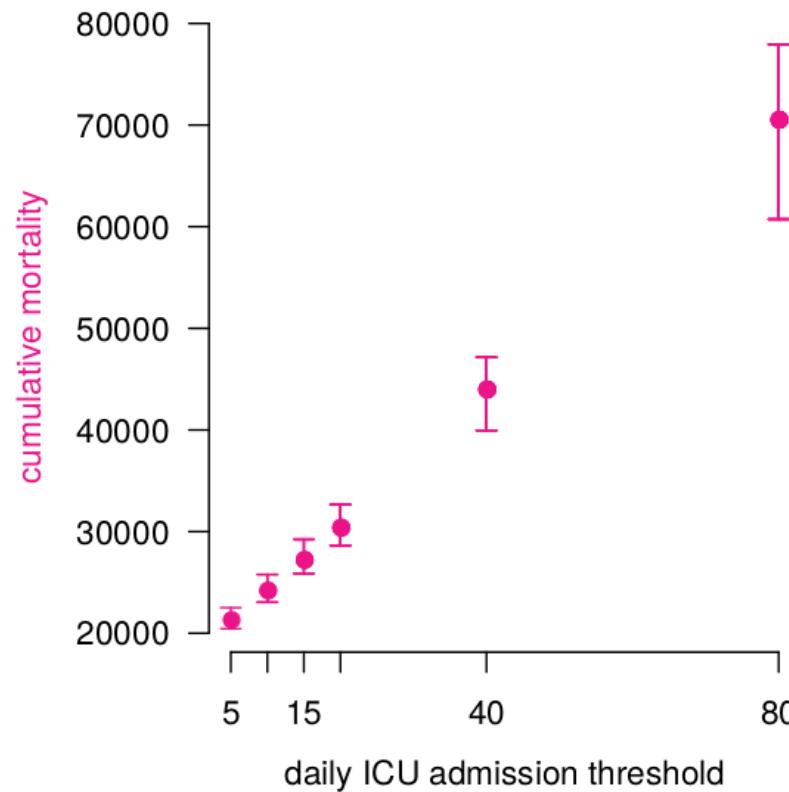


Figure S-8: **Adaptive lock-down threshold impact**

Each dot represents a simulation of the model with adaptive lock-down implemented from May 12. The abscissa show the lock-down triggering threshold in terms of daily ICU admissions nationwide. The ordinate represents the median final death toll by the end of the year, along with their 95%-confidence intervals. Lock-down lifting threshold variation has negligible impact on the results (not shown here).